

Applied Survival Analysis

Lab 6: Model Selection in Survival Analysis

Today, we are going to understand Collet's approach for model selection within the context of a proportional hazards model and to assess the overall fit of the model by checking the residuals.

1. Collet's Approach for Model Selection:

We are going to work with the MAC dataset (*mac.dta*), focusing on the outcome `dthstat` which equals 1 if a patient died and 0 otherwise. The time to death is `dthtime` and subjects who did not die are censored at their time of study discontinuation.

The covariates of interest for the purpose of this lab are:

<code>agecat</code>	<code>sex</code>	<code>cd4</code>	<code>karnof</code>
<code>ivdrug</code>	<code>antiret</code>	<code>rif</code>	<code>clari</code>

The significance of their effect will be tested using the Wald test. This time we are interested in the time to death. So we are going to *stset* the data in the following way:

```
stset dthtime, failure(dthstat)
```

Step 1: Fit univariate models to choose candidate predictors. Use criterion of $p \leq 0.15$ to identify predictors.

```
stcox agecat, nohr
```

```

      failure _d:  dthstat
analysis time _t: dthtime

Iteration 0:  log likelihood = -3393.2516
Iteration 1:  log likelihood = -3385.2229
Iteration 2:  log likelihood = -3385.2133
Refining estimates:
Iteration 0:  log likelihood = -3385.2133

Cox regression -- Breslow method for ties

No. of subjects =          1175          Number of obs   =          1175
No. of failures =           514
Time at risk    =          619081
Log likelihood  = -3385.2133          LR chi2(1)       =          16.08
                                          Prob > chi2    =          0.0001

```

_t	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
agecat	.3663803	.0928965	3.944	0.000	.1843065 .5484541

stcox sex, nohr

```

      failure _d:  dthstat
      analysis time _t:  dthtime

Iteration 0:   log likelihood = -3393.2516
Iteration 1:   log likelihood = -3392.0249
Iteration 2:   log likelihood = -3392.0109
Iteration 3:   log likelihood = -3392.0109
Refining estimates:
Iteration 0:   log likelihood = -3392.0109

Cox regression -- Breslow method for ties

No. of subjects =          1175          Number of obs   =          1175
No. of failures =           514
Time at risk    =        619081
Log likelihood   =   -3392.0109          LR chi2(1)       =          2.48
                                          Prob > chi2     =          0.1152
-----
      _t |
      _d |          Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      sex |   .2360791   .1453047    1.625  0.104   - .048713   .5208711
-----+-----

```

(a) Fit all other univariate models of interest and fill in the below summary table of univariate predictors:

<i>Predictor</i>	<i>Estimate</i>	<i>s.e.</i>	<i>p-value</i>	<i>HR</i>
agecat				
sex				
cd4				
karnof				
ivdrug				
antiret				
rif				
clari				

Is the effect of therapy with rif and/or clari significant?

Step 2:

- (i) Fit a multivariate model with all significant predictors ($p \leq 0.15$) from Step 1:
- (ii) Then use backwards selection to eliminate non-significant ones in a multivariate framework (use $p \leq 0.10$ for determining which ones to eliminate).
(To use automatic variable selection in STATA, we use the `sw stcox` command).

```
sw stcox agecat sex cd4 karnof antiret, pr(0.10)
```

```
(2 obs. dropped due to estimability)
LR test          begin with full model
p < 0.1000      for all terms in model
```

```
Cox regression -- Breslow method for ties
```

```
No. of subjects =          1175          Number of obs   =          1175
No. of failures =           514
Time at risk    =          619081
Log likelihood  = -3318.3627          LR chi2(5)       =          149.78
                                          Prob > chi2      =           0.0000
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
agecat	1.420866	.1334364	3.74	0.000	1.181993 1.708013
sex	1.369403	.2001345	2.15	0.031	1.028326 1.823611
cd4	.9895408	.001542	-6.75	0.000	.9865232 .9925676
karnof	.9626493	.0049101	-7.46	0.000	.9530736 .9723211
antiret	.7926794	.0786474	-2.34	0.019	.652595 .9628339

Step 3:

Use forwards selection to add any variables not significant at Step 1 to the multivariate model obtained at the end of Step 2. To be conservative, use $p \leq 0.10$ for deciding whether to add variables or not.

Note that we specify the option `lockterm1` to force the variables in the first parenthesis into the model.

```
sw stcox (agecat sex cd4 karnof antiret) ivdrug (rif clari), pe(0.10)
lockterm1
```

```
(2 obs. dropped due to estimability)
begin with term 1 model
p >=0.1000      for all terms in model
```

```
Cox regression -- Breslow method for ties
```

```
No. of subjects =          1175          Number of obs   =          1175
No. of failures =           514
Time at risk    =          619081
Log likelihood  = -3318.3627          LR chi2(5)       =          149.78
                                          Prob > chi2      =           0.0000
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
agecat	1.420866	.1334364	3.74	0.000	1.181993 1.708013
sex	1.369403	.2001345	2.15	0.031	1.028326 1.823611
cd4	.9895408	.001542	-6.75	0.000	.9865232 .9925676
karnof	.9626493	.0049101	-7.46	0.000	.9530736 .9723211
antiret	.7926794	.0786474	-2.34	0.019	.652595 .9628339

(b) Are there any other variables added to the model?

Step 4:

Create all possible 2-way interaction terms based on the main effects in the model at the end of Step 3. Add these to the multivariate model and use a backwards selection procedure to eliminate those not significant at $p = 0.10$. Remember to force the inclusion of all of the main effects from the end of Step 3 in the model.

First we are going to generate all possible 2-way interactions:

```

gen agsex=agecat*sex
gen agcd4=agecat*cd4
gen agkar=agecat*karnof
gen aganti=agecat*antiret
gen sexcd4=sex*cd4
gen sexkar=sex*karnof
gen sexanti=sex*antiret
gen cd4kar=cd4*karnof
gen cd4anti=cd4*antiret
gen karanti=karnof*antiret

```

(c) Now fit the appropriate model that was described above in Step 4.

Step 5:

Check if all variables are significant. If a main effect has become non-significant and there are no interactions involving this main effect in the model at the end of Step 4, then you may consider excluding it. Discretion is needed in determining whether to include covariates and/or interactions that are marginally significant. At this stage, for the purposes of this exercise, use a somewhat stricter significance level of $\alpha = 0.02$ to account for the multiple tests we have conducted.

```

sw stcox agecat sex cd4 karnof antiret sexanti karanti , pe(0.02)
(2 obs. dropped due to estimability)
                                begin with empty model
p = 0.0000 < 0.0200  adding  karnof
p = 0.0000 < 0.0200  adding  cd4
p = 0.0002 < 0.0200  adding  agecat

Cox regression -- Breslow method for ties

No. of subjects =          1175                Number of obs   =          1175
No. of failures =           514
Time at risk    =          619081

Log likelihood   =   -3322.7981                LR chi2(3)       =          140.91
                                                Prob > chi2     =           0.0000

-----+-----
      _t | Haz. Ratio   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      karnof |   .9628398   .0049084    -7.43   0.000   .9532673   .9725083
      cd4    |   .9895551   .0015446    -6.73   0.000   .9865324   .9925871
      agecat |   1.420961   .1330381     3.75   0.000   1.182736   1.707168
-----+-----

```

Step 6:

You may also want to include some models with and without certain marginally significant covariates to evaluate the change in the AIC criterion or consider alternate codings of covariates (eg. cd4cat instead of cd4 or age instead of agecat).

This time we could use the stepwise backwards procedure with probability of entry $pe = 0.05$ and probability of removal $pr = 0.10$.

(i) (cd4cat instead of cd4)

```
sw stcox agecat sex cd4cat karnof antiret, pe(0.05) pr(0.10)
```

```
(2 obs. dropped due to estimability)
```

```
begin with full model  
p < 0.1000 for all terms in model
```

```
Cox regression -- Breslow method for ties
```

```
No. of subjects = 1175 Number of obs = 1175  
No. of failures = 514  
Time at risk = 619081  
Log likelihood = -3326.9922 LR chi2(5) = 132.52  
Prob > chi2 = 0.0000
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
agecat	1.414087	.1329843	3.68	0.000	1.176053 1.700299
sex	1.390386	.2033177	2.25	0.024	1.04391 1.851857
cd4cat	.5762558	.0525747	-6.04	0.000	.4818989 .689088
karnof	.9627031	.0049187	-7.44	0.000	.9531107 .972392
antiret	.7910011	.0785034	-2.36	0.018	.651177 .9608491

(ii) (age instead of agecat)

In this case we are including age instead of agecat. To use the stepwise forwards procedure we have to add the option **forward** (the default is backwards):

```
sw stcox age sex cd4 karnof antiret, pe(0.05) pr(0.10) forward
```

```
begin with empty model  
p = 0.0000 < 0.0500 adding karnof  
p = 0.0000 < 0.0500 adding cd4  
p = 0.0000 < 0.0500 adding age  
p = 0.0406 < 0.0500 adding sex  
p = 0.0351 < 0.0500 adding antiret
```

```
Cox regression -- Breslow method for ties
```

```
Entry time 0 Number of obs = 1177  
LR chi2(5) = 153.07  
Prob > chi2 = 0.0000  
Log likelihood = -3316.7163 Pseudo R2 = 0.0226
```

dthtime	dthstat	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
karnof		-.0376065	.0051224	-7.342	0.000	-.0476463 -.0275666
cd4		-.0105009	.0015625	-6.720	0.000	-.0135635 -.0074384
age		.0211124	.0049501	4.265	0.000	.0114104 .0308145
sex		.3210042	.1460386	2.198	0.028	.0347737 .6072346
antiret		-.2099265	.0996185	-2.107	0.035	-.4051752 -.0146778

(d) Summarize the final models at the ends of Steps 1-6 in a table, such as that shown below. Use $\alpha = 3$ in calculating the AIC value ($AIC = -2\log L + (\alpha * q)$) for each of the models below. Which model appears best in terms of the AIC criterion?

<i>Model</i>	<i>Covariates</i>	<i>-2logL</i>	<i>q</i>	<i>AIC</i>
Step 2 (i)	agecat, karnof, sex, cd4, antiret	6636.7	5	
Step 2 (ii)	(same as above)			
Step 3	(same as above)			
Step 4	Step3 +karnof*antiret,sex*antiret	6628.6	7	
Step 5	agecat, karnof, cd4	6645.6	3	
Step 6 (i)	agecat, karnof, sex, cd4cat, antiret	6654.0	5	
Step 6 (ii)	age, karnof, sex, cd4, antiret	6633.4	5	

2. Assessing Overall Model Fit:

We will assess the overall fit of the model in Step 6 (ii) .

◆ Martingale Residuals:

First we will fit the model along with the option `mgale (newvar)` to get the *Martingale* residuals:

```
stcox age sex cd4 karnof antiret , mgale(mg)
```

```
      failure _d:  dthstat
analysis time _t:  dthtime
```

```
Iteration 0:  log likelihood = -3393.2516
Iteration 1:  log likelihood = -3318.6397
Iteration 2:  log likelihood = -3316.7201
Iteration 3:  log likelihood = -3316.7163
Refining estimates:
Iteration 0:  log likelihood = -3316.7163
```

Cox regression -- Breslow method for ties

```
No. of subjects =          1175          Number of obs   =          1175
No. of failures =           514
Time at risk   =          619081
Log likelihood  = -3316.7163          LR chi2(5)       =          153.07
                                          Prob > chi2     =           0.0000
```

```
-----+-----
```

<u>_t</u>	<u>_d</u>	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
age		1.021337	.0050557	4.265	0.000	1.011476	1.031294
sex		1.378511	.2013159	2.198	0.028	1.035385	1.835349
cd4		.989554	.0015462	-6.720	0.000	.9865281	.9925892
karnof		.9630919	.0049334	-7.342	0.000	.953471	.9728098
antiret		.8106438	.0807551	-2.107	0.035	.66686	.9854294

```
-----+-----
```

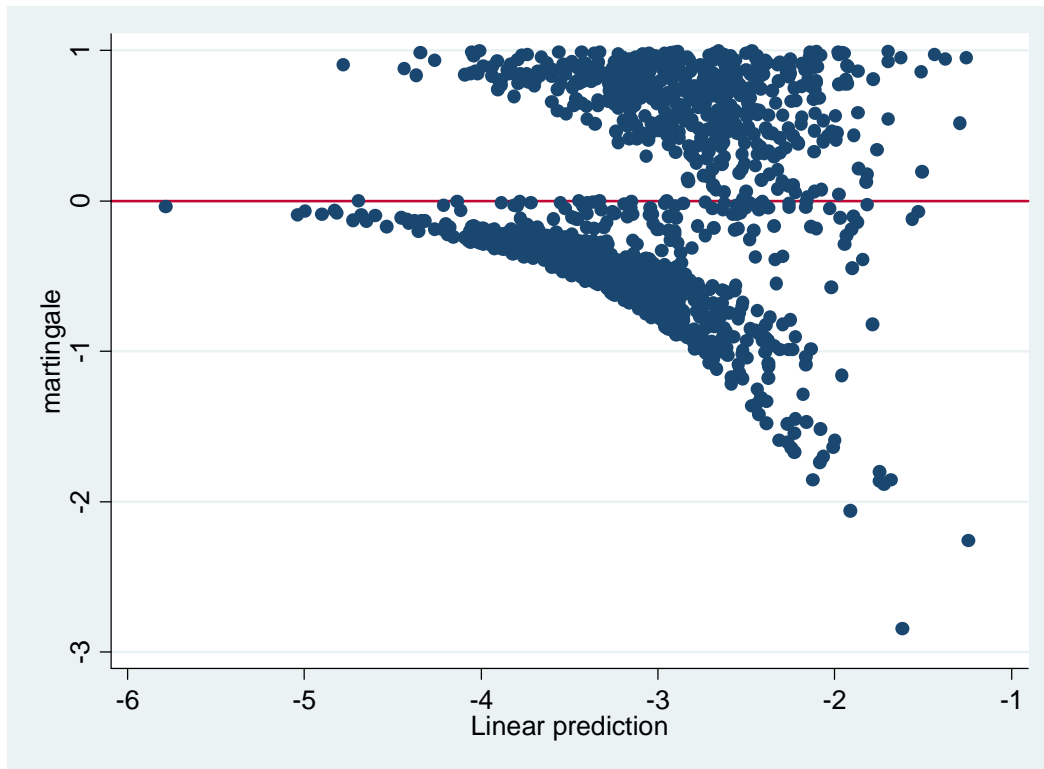
Once the martingale residuals are created, you usually plot them versus the predicted log HR or any of the individual covariates to assess the model fit.

To get the predicted log HR we use the following command:

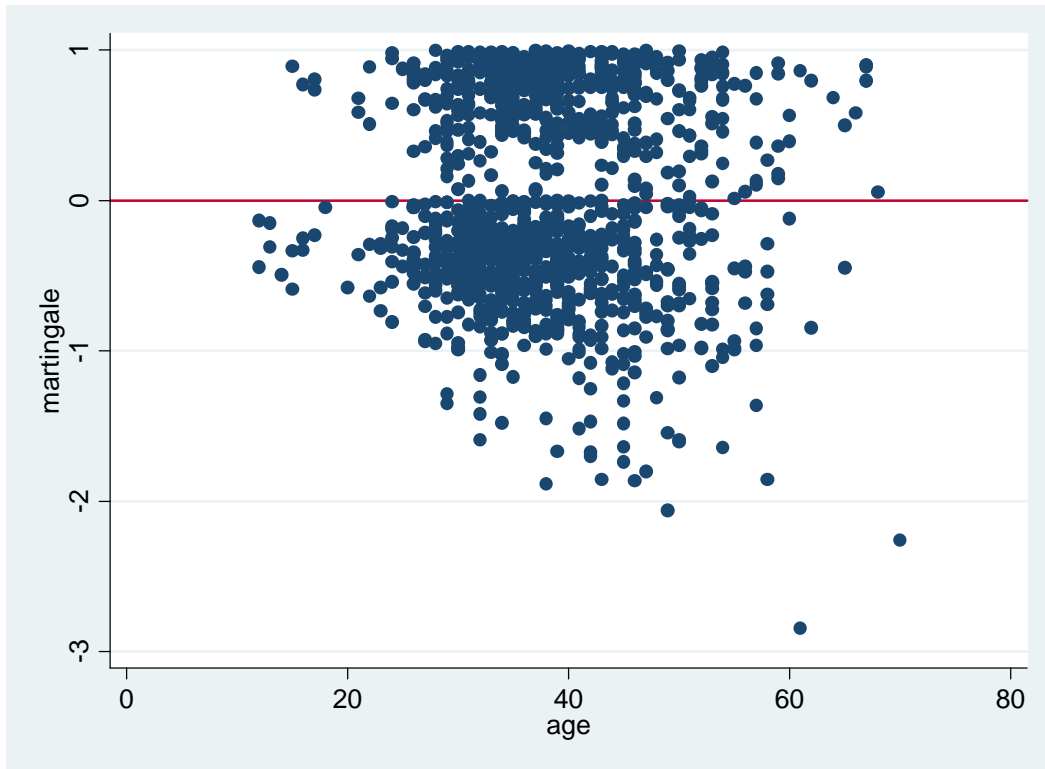
```
predict betaz, xb  
(2 missing values generated)
```

And the graphs:

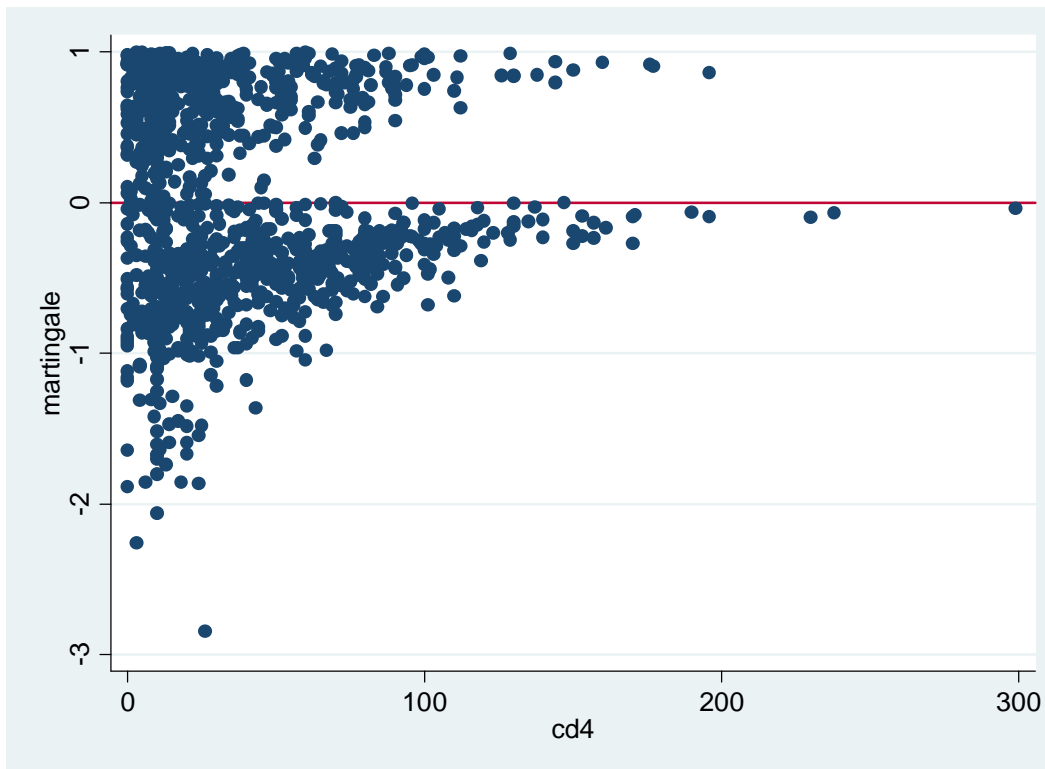
```
scatter mg betaz, yline(0)
```



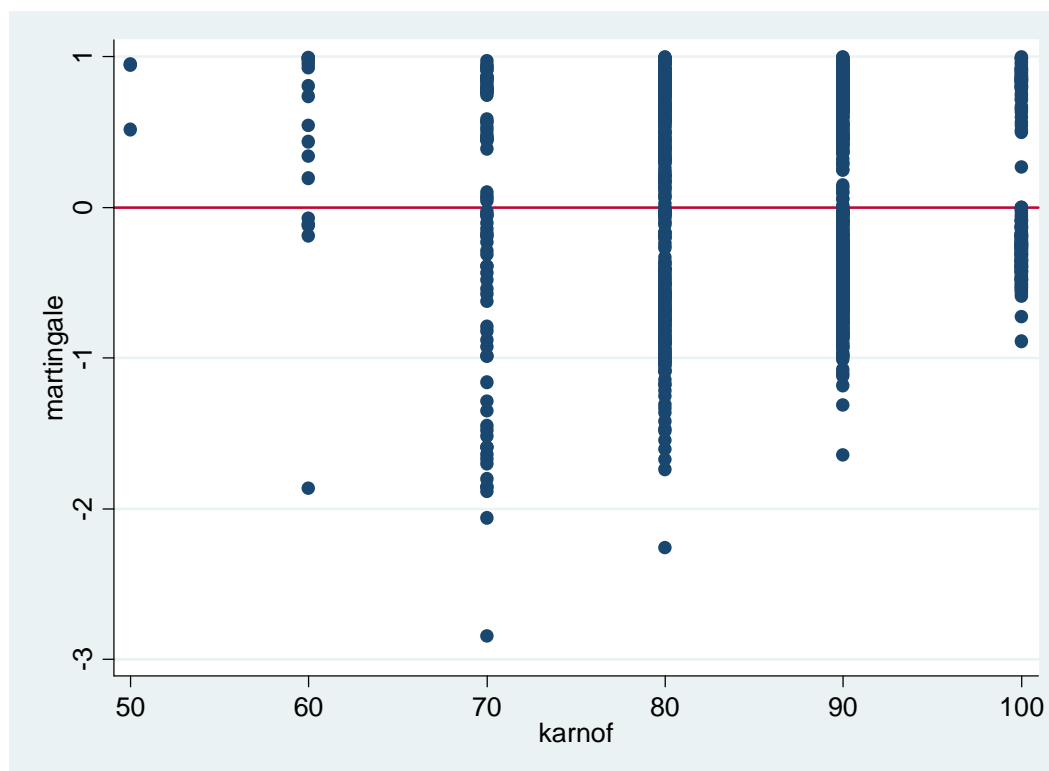
```
scatter mg age, yline(0)
```



```
scatter mg cd4, yline(0)
```




```
scatter mg karnof, yline(0)
```



◆ **Generalized (Cox-Snell) Residuals:**

To get generalized residuals in STATA we use the previous `stcox` command with the `mgale` option and then use the `predict` command with `csnell` option:

```
stcox age sex cd4 karnof antiret , mgale(mg)
```

```
predict csres, csnell
(2 missing values generated)
```

Then produce the informative graph for the generalised residuals we have to define a survival dataset using the Cox-Snell residuals as the “pseudo” failure times.

```
stset csres, failure(dthstat)
```

```
failure event: dthstat ~= 0 & dthstat ~= .
obs. time interval: (0, csres]
exit on or before: failure
```

```
-----
1177 total obs.
   2 event time missing (csres==.)          PROBABLE ERROR
   3 obs. end on or before enter()
-----
1172 obs. remaining, representing
  514 failures in single record/single failure data
  514 total analysis time at risk, at risk from t =          0
                                     earliest observed entry t =          0
                                     last observed exit t = 2.886069
```

```

sts gen survcs=s

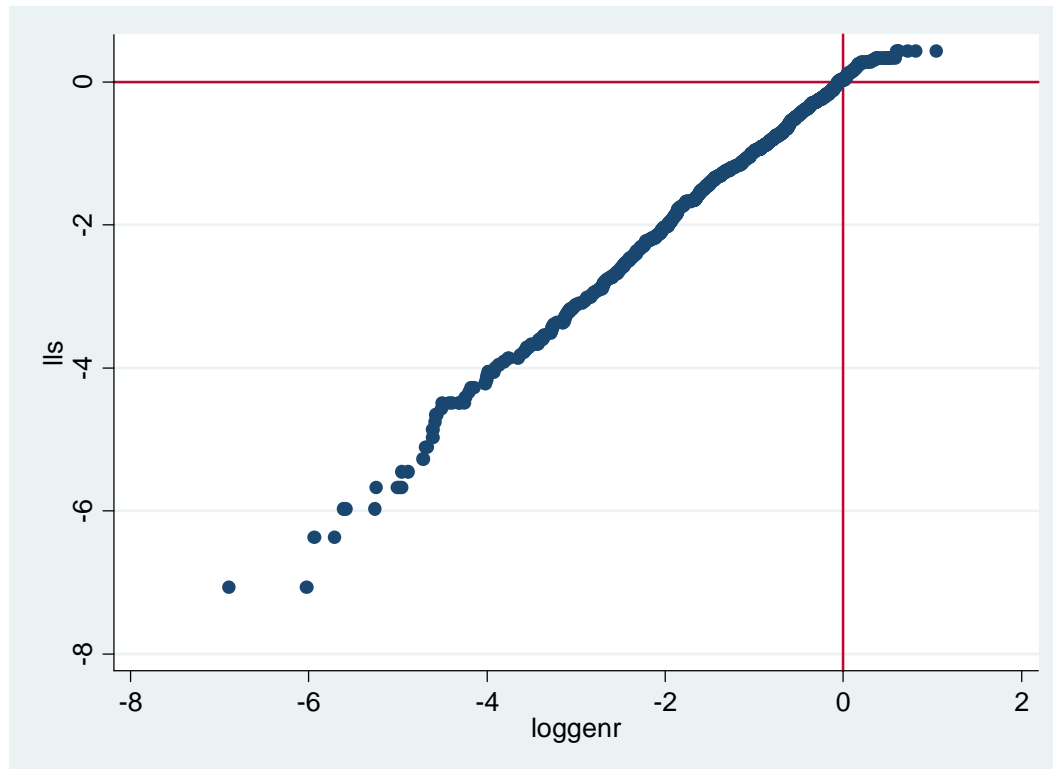
gen lls=log(-log(survcs))
(6 missing values generated)

gen loggenr=log(csres)
(5 missing values generated)

```

Then we want to plot :

```
scatter lls loggenr, yline(0) xline(0)
```



If the data fit the model well we would expect a straight line.

◆ Deviance Residuals:

Again to get the deviance residuals in STATA we first use the `stcox` command with the `mgale` option (make sure to drop `mg` before and `stset` the dataset using the real failure times) and then the `predict` command with the `deviance` option this time.

```

drop mg
stset dthtime dthstat

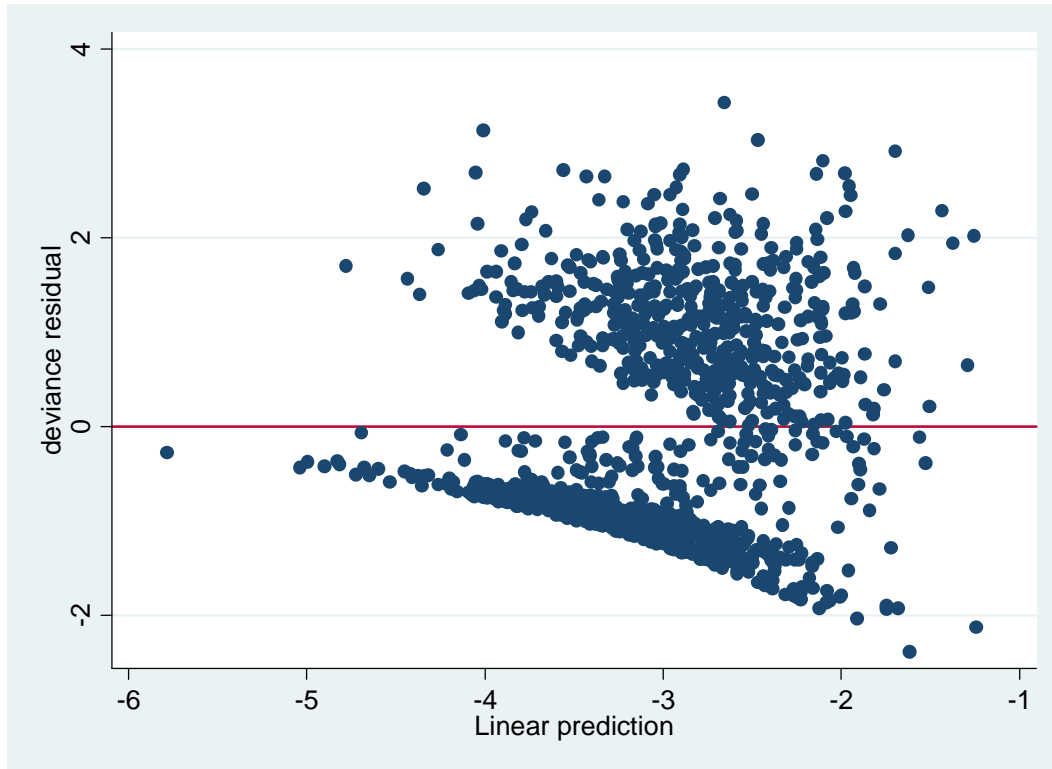
stcox age sex cd4 karnof antiret , mgale(mg)

predict devres, deviance
(5 missing values generated)

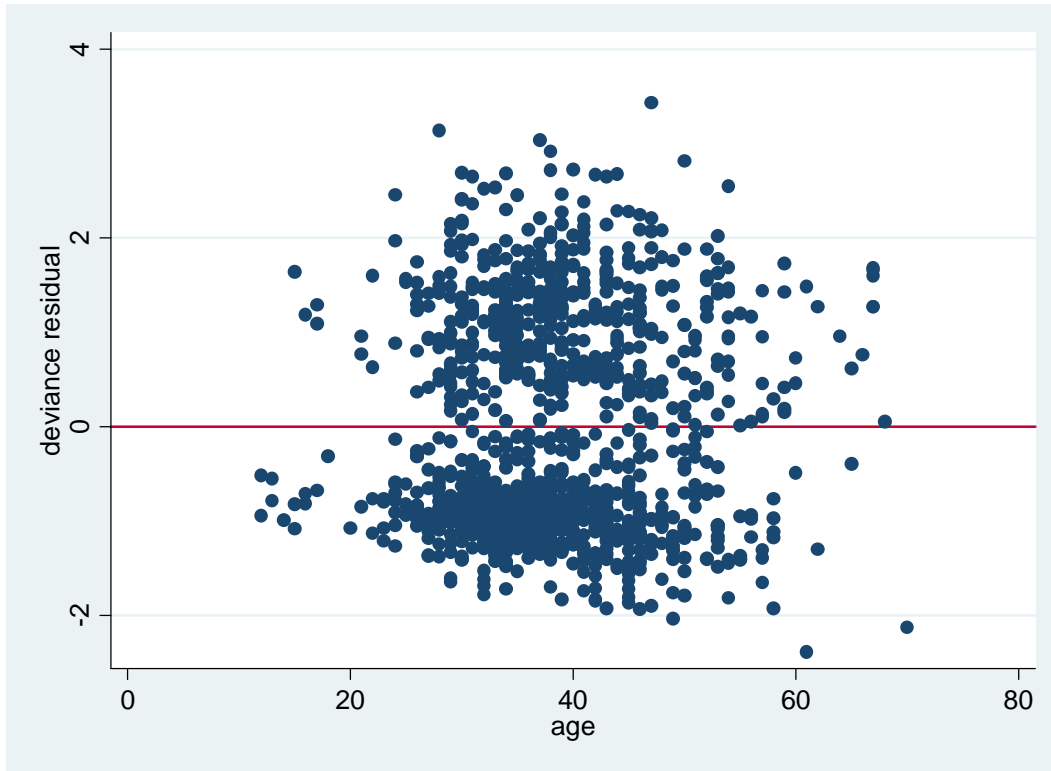
```

And they can be plotted against the predicted log(HR) and other covariates, as shown for the Martingale residuals.

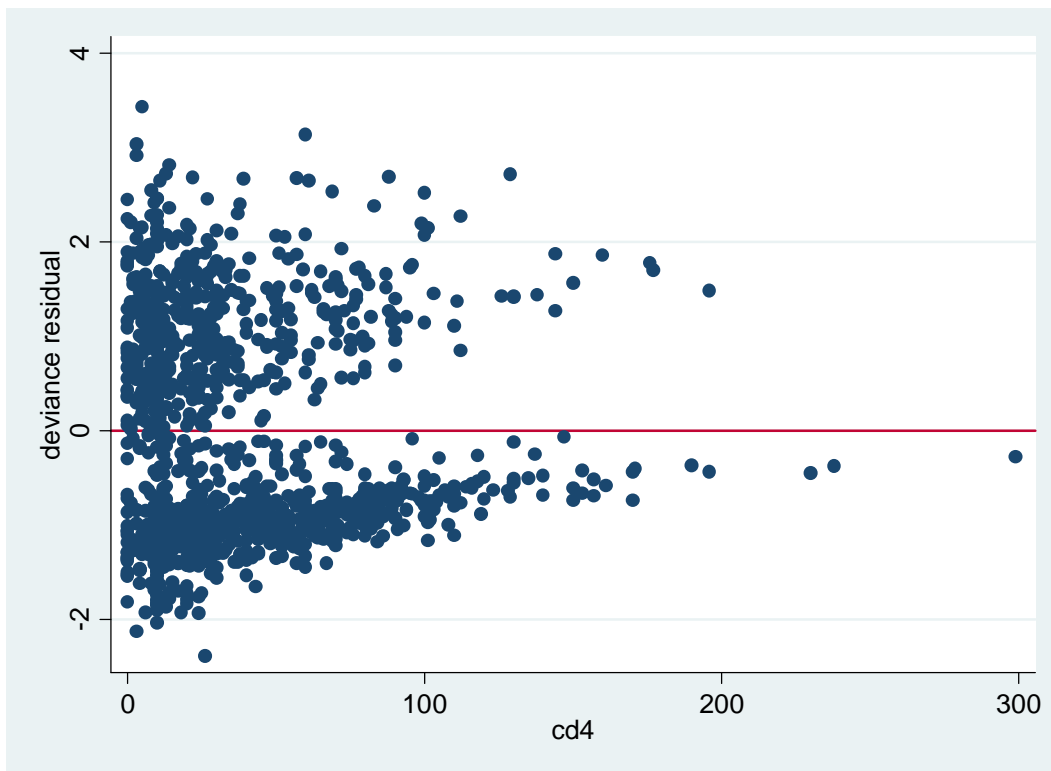
```
scatter devres betaz, yline(0)
```



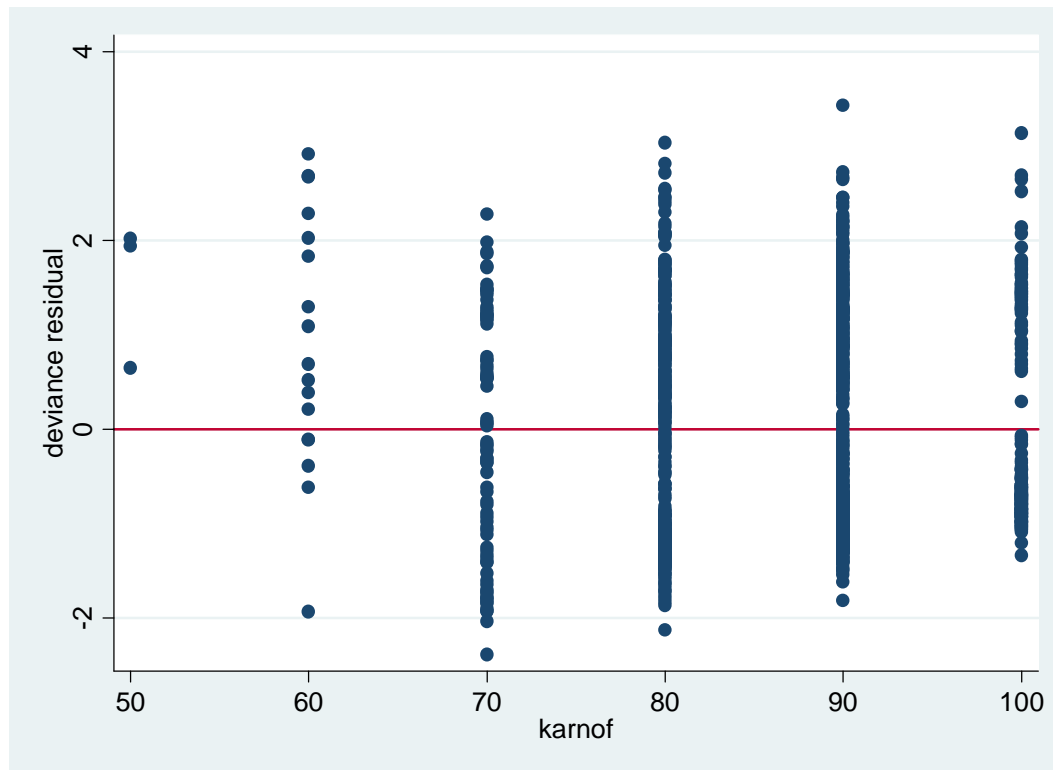
```
scatter devres age, yline(0)
```



```
scatter devres cd4, yline(0)
```



```
scatter devres karnof, yline(0)
```



◆ **Schoenfeld Residuals:**

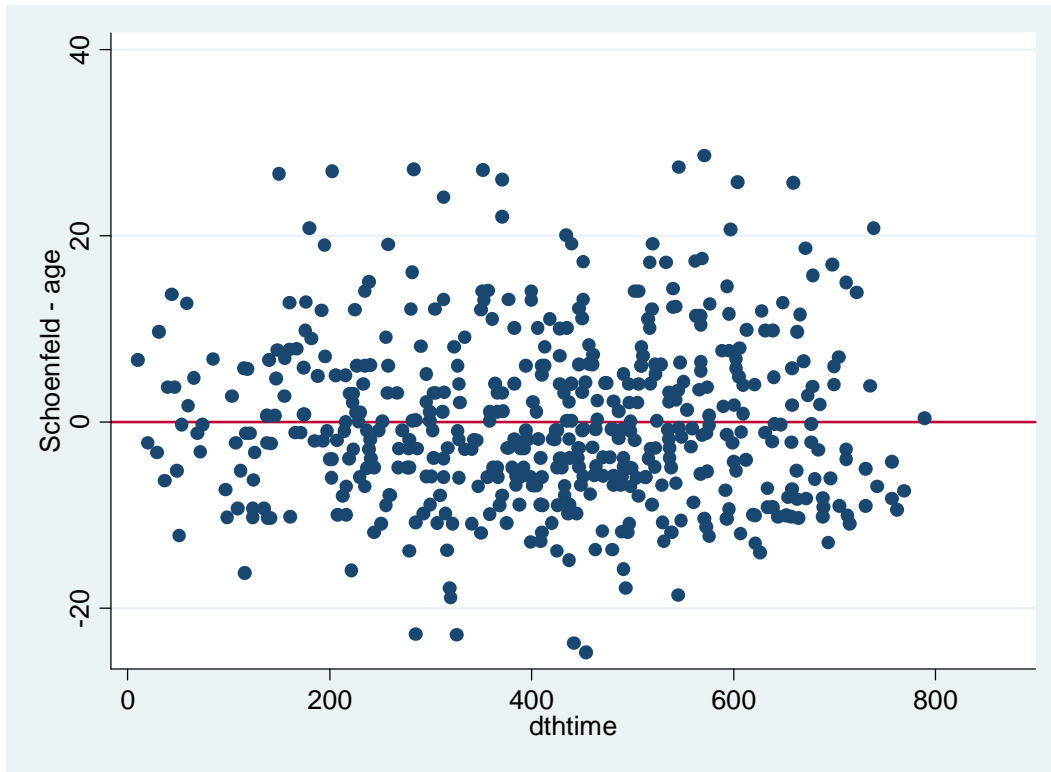
In STATA the *Schoenfeld* residuals are generated in the `stcox` command itself, using the `schoenfeld (newvar(s))` option:

```
stcox age sex cd4 karnof antiret , schoenfeld(ageres sexres cd4res  
karnres antires)
```

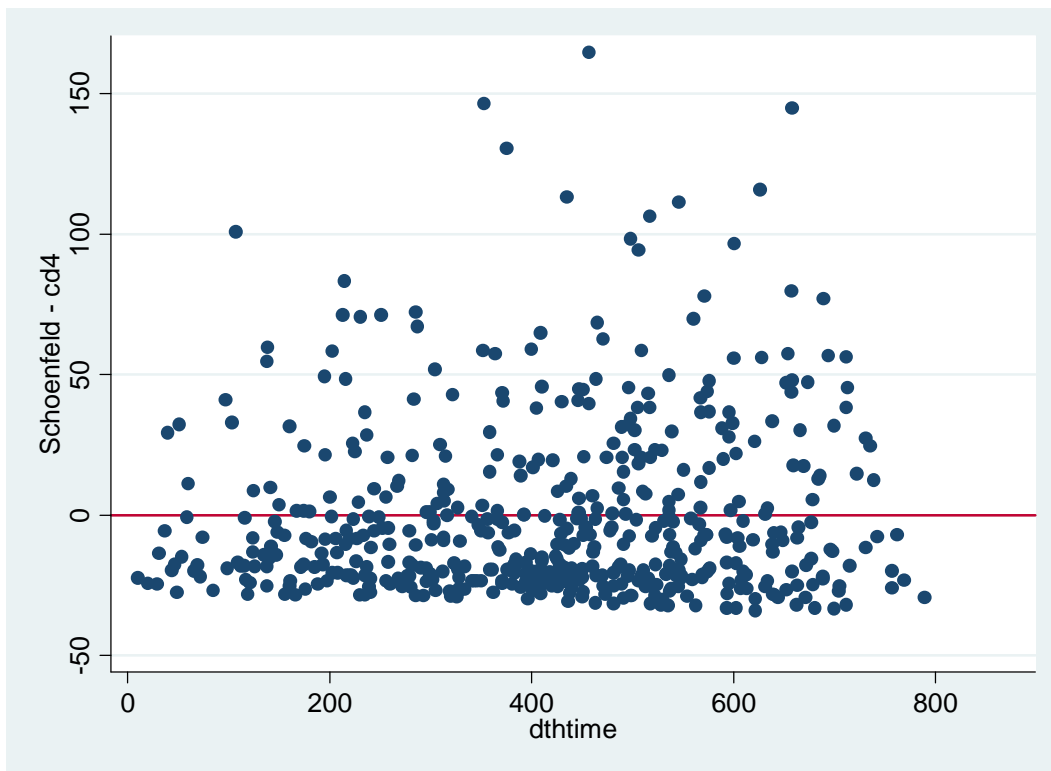
(The output is exactly the same as before)

Then we plot them against event time:

```
scatter ageres dthtim, yline(0)
```



```
scatter cd4res dthtim, yline(0)
```



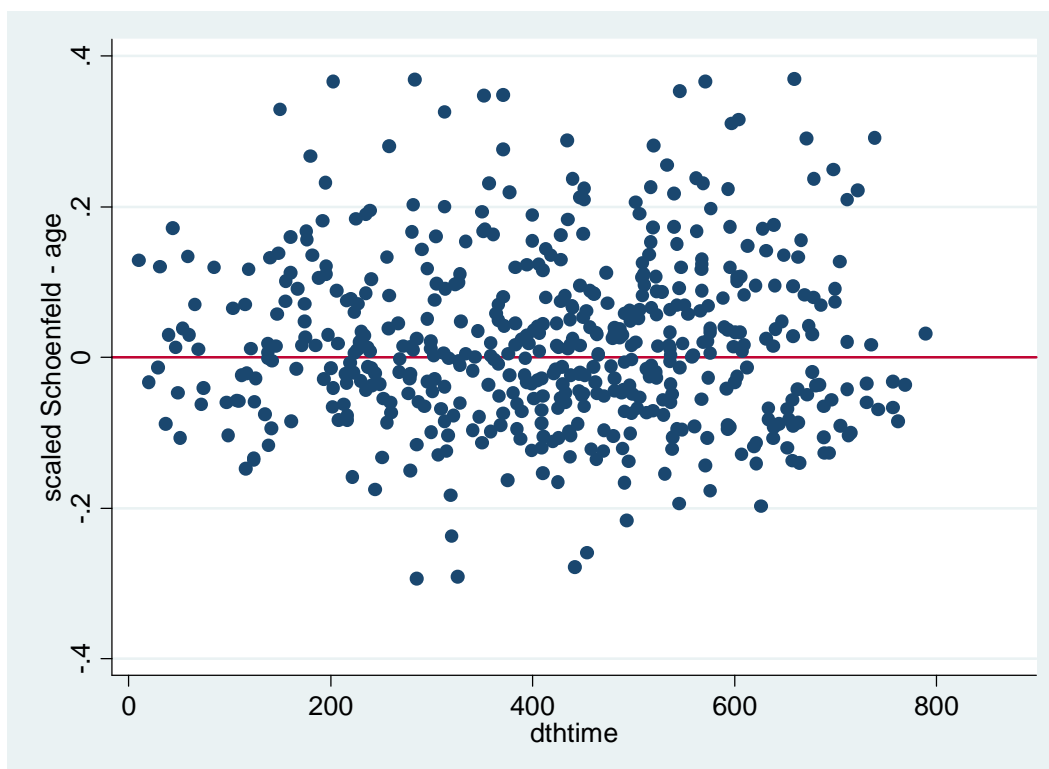
◆ **Weighted Schoenfeld Residuals:**

These residuals are used more than the previous unweighted version, because they are symmetric around 0. In this case we use the following command:

```
stcox age sex cd4 karnof antiret , scaledsch(ageres2 sexres2 cd4res2  
karnres2 antires2)
```

(Same output as before)

```
scatter ageres2 dthtim, yline(0)
```



```
scatter cd4res2 dthtim, yline(0)
```

