

# Strong Law of Large Numbers (SLLN) ↓

Theorem / Ισχυρός Νόμος των Μεγάλων Αριθμίων (INMA)

Let  $(X_n)_{n \geq 1}$  be a sequence of i.i.d. r.v.'s on  $(\Omega, \mathcal{A}, P)$

with  $X_n \in \overline{\mathbb{R}}, n \geq 1$  and  $E|X_1| < +\infty$ . Set  $\mu = E(X_1)$

and  $S_n = \sum_{k=1}^n X_k, \forall n \geq 1$ . Then  $\overline{X}_n = \frac{S_n}{n} \xrightarrow{\text{a.s.}} \mu$ .

Proof: We will assume that  $E(X_1^2) < +\infty$  (to simplify the proof)

$E(X_1^2) < +\infty \Rightarrow V(X_1) < +\infty$ . We set  $\sigma^2 = V(X_1)$ .

S1 (step 1) We prove the theorem for  $X_n \geq 0$ .

Set  $Y_n = \overline{X}_n - \mu, n \geq 1$

$$E(Y_n) = E(\overline{X}_n - \mu) = E(\overline{X}_n) - \mu = E(X_1) - \mu = \mu - \mu = 0$$

$$E(Y_n^2) = V(Y_n) = V(\overline{X}_n) = \frac{V(X_1)}{n} = \frac{\sigma^2}{n}, \sigma^2 < +\infty$$

Take the subsequence  $(Y_{n^2})_{n \geq 1}$  of  $(Y_n)_{n \geq 1}$ . 2.

$$E\left(\sum_{n \geq 1} Y_{n^2}^2\right) \stackrel{B-L}{=} \sum_{n \geq 1} E(Y_{n^2}^2) = \sum_{n \geq 1} \frac{\sigma^2}{n^2} = \sigma^2 \sum_{n \geq 1} \frac{1}{n^2} < +\infty$$

$$\Rightarrow \sum_{n \geq 1} Y_{n^2}^2 < +\infty \text{ with prob. 1.}$$

$$\Rightarrow Y_{n^2}^2 \xrightarrow{\text{a.s.}} 0 \stackrel{\text{CMT}}{\Rightarrow} |Y_{n^2}| \xrightarrow{\text{a.s.}} 0 \Rightarrow Y_{n^2} \xrightarrow{\text{a.s.}} 0$$

$$\Rightarrow \frac{\sum_{n^2 \leq k} Y_{n^2}^2}{n^2} - \mu \xrightarrow{\text{a.s.}} 0 \Rightarrow \boxed{\frac{\sum_{n^2 \leq k} Y_{n^2}^2}{n^2} \xrightarrow{\text{a.s.}} \mu} \quad (1)$$

$\forall k \geq 1$ , set  $n(k) = \lfloor \sqrt{k} \rfloor$

But,  $\lfloor \sqrt{k} \rfloor \leq \sqrt{k} \leq \lfloor \sqrt{k} \rfloor + 1 \Rightarrow n(k) \leq \sqrt{k} \leq n(k) + 1$

$$\Rightarrow \frac{1}{n(k)+1} \leq \frac{1}{\sqrt{k}} \leq \frac{1}{n(k)} \Rightarrow \boxed{\frac{1}{(n(k)+1)^2} \leq \frac{1}{k} \leq \frac{1}{n^2(k)}} \quad (2)$$

Since  $n^2(k) \leq k \leq (n(k)+1)^2$ , we have

$$\boxed{\sum_{n^2(k)} \leq \sum_k \leq \sum_{(n(k)+1)^2}} \quad (3)$$

$$\boxed{\frac{\sum_{n^2(k)} \leq \sum_k \leq \sum_{(n(k)+1)^2}}{(n(k)+1)^2} \leq \frac{\sum_k}{k} \leq \frac{\sum_{(n(k)+1)^2}}{n^2(k)}} \quad (4)$$

But

$$\cdot \frac{S_{n^2(k)}}{(n(k)+1)^2} = \frac{S_{n^2(k)}}{n^2(k)} \cdot \left( \frac{n(k)}{n(k)+1} \right)^2 \xrightarrow[k \rightarrow \infty]{\text{a.s.}} \mu \cdot 1 = \mu$$

( $n(k) \xrightarrow[k \rightarrow \infty]{} \infty$ )

$$\cdot \frac{S_{(n(k)+1)^2}}{n^2(k)} = \frac{S_{(n(k)+1)^2}}{(n(k)+1)^2} \cdot \left( \frac{(n(k)+1)^2}{n^2(k)} \right) \xrightarrow{\text{a.s.}} \mu \cdot 1 = \mu$$

By sandwich  $\frac{S_k}{k} \xrightarrow[k \rightarrow \infty]{\text{a.s.}} \mu$

$$\underline{S2} \mid X_n \in \overline{\mathbb{R}} \parallel X_i = X_i^+ - X_i^-$$

$$\frac{S_n}{n} = \frac{X_1 + \dots + X_n}{n} = \frac{X_1^+ + \dots + X_n^+}{n} - \frac{X_1^- + \dots + X_n^-}{n}$$

$X_n$  i.i.d  $\Rightarrow (X_n^+)_{n \geq 1}$  and  $(X_n^-)_{n \geq 1}$  are i.i.d. 4.

Also

$$\left. \begin{aligned} X_1^+ = \max\{0, X_1\} \leq |X_1| \\ X_1^- = \max\{0, -X_1\} \leq |X_1| \end{aligned} \right\} \Rightarrow \left. \begin{aligned} (X_1^+)^2 \leq X_1^2 \\ (X_1^-)^2 \leq X_1^2 \end{aligned} \right\} \Rightarrow \begin{aligned} E(X_1^{+2}) < +\infty \\ E(X_1^{-2}) < +\infty \end{aligned}$$

So, the conditions of SL hold (i.i.d + second moment finite)

$$\frac{X_1^+ + \dots + X_n^+}{n} \xrightarrow{\text{a.s.}} E(X_1^+) \in \mathbb{R}$$

$$\frac{X_1^- + \dots + X_n^-}{n} \xrightarrow{\text{a.s.}} E(X_1^-) \in \mathbb{R}$$

$$\left. \begin{aligned} \frac{X_1^+ + \dots + X_n^+}{n} \xrightarrow{\text{a.s.}} E(X_1^+) \in \mathbb{R} \\ \frac{X_1^- + \dots + X_n^-}{n} \xrightarrow{\text{a.s.}} E(X_1^-) \in \mathbb{R} \end{aligned} \right\} \Rightarrow \frac{X_1^+ + \dots + X_n^+ - X_1^- - \dots - X_n^-}{n} \xrightarrow{\text{a.s.}} E(X_1^+) - E(X_1^-) \\ \parallel \\ E(X_1)$$

Rem The theorem also holds for  $E(X_1) = +\infty$   
or  $E(X_1) = -\infty$ . Only exception if  $E(X_1^+) = E(X_1^-) = +\infty$

## Εφαρμογή (Μέθοδος Ογκλήρωσης Monte-Carlo)

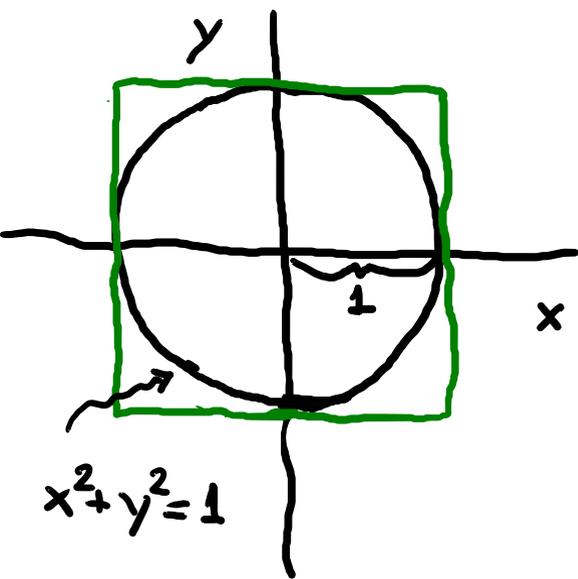
5.

Είναι μια μέθοδος προσέγγισης ογκλήρωμάτων, δηλ. της τιμής ενός ογκλήρωματος, που μπορεί να εκφραστεί ως μέση τιμή κάποιας τ.μ.

Προϋποθέτει τη δυνατότητα προσομοίωσης από την κατανομή αυτής της τ.μ.

Παράδειγμα : προσεγγιστικός υπολογισμός του  $\pi$

$$\pi \approx 3,1415 \dots$$



$$E_{\text{prob.}} = \pi \cdot r^2 \stackrel{r=1}{=} \pi \quad 6.$$

$$\pi = \iint_{x^2+y^2 \leq 1} 1 \cdot dx dy = \iint_{\substack{-1 \leq x \leq 1 \\ -1 \leq y \leq 1}} \frac{1}{\{x^2+y^2 \leq 1\}} dx dy$$

- Idea:  $(X, Y)$  such that  $\left. \begin{array}{l} X, Y \sim \text{Unif}(-1, 1) \\ X \text{ indep. of } Y. \end{array} \right\} f(x,y) = \frac{1}{4} \quad (|x|, |y| \leq 1)$

Rewrite  $\pi = 4 \cdot \frac{\pi}{4} = 4 \cdot \int_{-1}^1 \int_{-1}^1 \frac{1}{\{x^2+y^2 \leq 1\}} \cdot \frac{1}{4} dx dy$   
 $= 4 E \left[ \frac{1}{\text{Disk}} \cdot (X, Y) \right], \text{Disk} = \{(x,y) \in \mathbb{R}^2 : x^2+y^2 \leq 1\}$

$$\Rightarrow \frac{\pi}{4} = E \left[ \underbrace{1_{\text{Disk}}(x, y)}_{g(x, y)} \right]$$

7.

Approximation via Monte-Carlo

Draw  $X_1, X_2, \dots, X_n \sim \text{Unif}(-1, 1)$ , all i.i.d.  
 $Y_1, Y_2, \dots, Y_n \sim \text{Unif}(-1, 1)$  (from your favorite simulator)

Then  $Z_1 = 1_{\text{Disk}}(X_1, Y_1), Z_2 = 1_{\text{Disk}}(X_2, Y_2), \dots, Z_n = 1_{\text{Disk}}(X_n, Y_n), \dots$   
 are i.i.d.,  $Z_n \sim \text{Be}(p)$ ,  $p = \underbrace{P(Z_n=1)}_{P(X_n^2 + Y_n^2 \leq 1)} = \frac{\pi}{4}$

Take  $\bar{Z}_n = \frac{Z_1 + \dots + Z_n}{n}$   $\xrightarrow{\text{a.s.}} \frac{\pi}{4}$   $\frac{E[1_{\text{Disk}}(Z_n)]}{1}$   
 percentage of points inside disk  $\quad \quad \quad \text{SLLN} \quad \quad \quad \text{for large } n: \pi \approx 4 \cdot \bar{Z}_n$