

Probabilities II / Τετάρτη

Covariance of 2 random variables

Definition: Let $X, Y \in \mathcal{L}^2(P)$ and $E(X, Y)$ well defined. Then the quantity

$$C(X, Y) \text{ or } \text{Cov}(X, Y) = E \left[\underbrace{(X - E(X))}_{\in \mathbb{R}} \underbrace{(Y - E(Y))}_{\in \mathbb{R}} \right] \in \mathbb{R}$$

is said to be the covariance of X and Y . (συνδιακύμανση / συνδιαστροφή)

Properties

$$(1) C(X, Y) = E(XY) - E(X) \cdot E(Y)$$

$$(2) C(aX + \beta, \gamma Y + \delta) = a\gamma \cdot C(X, Y)$$

$$(3) C(X+Y, Z+W) = C(X, Y) + C(X, W) + C(Y, Z) + C(Y, W)$$

Proposition

$$X, Y \in \mathcal{L}^2(P) \rightarrow |C(X, Y)| \leq \sqrt{V(X)} \sqrt{V(Y)}$$

Proof

$$|C(X, Y)| = \left| E \left[\underbrace{(X - \mu_X)}_{\tilde{X}} \underbrace{(Y - \mu_Y)}_{\tilde{Y}} \right] \right| \stackrel{C-S}{\leq} \left(E(\tilde{X}^2) \right)^{1/2} \left(E(\tilde{Y}^2) \right)^{1/2} \\ \underbrace{\qquad\qquad\qquad}_{V(X)} \qquad\qquad\qquad \underbrace{\qquad\qquad\qquad}_{V(Y)}$$

Modes of Convergence of sequences of random variables (Τρόποι σύγκλισης)

$$i) X_n \xrightarrow{a.s.} X \stackrel{\text{def}}{\iff} P(\lim X_n = X) = 1, \text{ where}$$

$$\{\lim X_n = X\} = \{\omega \in \Omega : \exists \lim X_n(\omega) \text{ και } \lim X_n(\omega) = X(\omega)\}$$

$$X_n \xrightarrow{L^p} X \iff \forall p \in (0, \infty), E(|X_n - X|^p) \xrightarrow{n \rightarrow \infty} 0 \quad \text{or} \quad (E(|X_n - X|^p))^{1/p} \rightarrow 0 \iff E(|X_n - X|^p) \rightarrow 0$$

($d_p(X_n, X) \rightarrow 0$) or $\forall p \in (0, \infty], \|X_n - X\|_p \rightarrow 0$
 ($p \in (0, 1)$ δεν είναι νόρμα)

Για $p=0, E|X_n - X|^0 = P(X_n \neq X), \|X_n - X\|_p = (E|X_n - X|^p)^{1/p}$

(we had defined ~~ess sup~~ $\|X_n - X\|_p = \text{ess sup } |X_n - X|$)

$$(\|X\|_\infty = \text{ess sup } |X| = \inf \{M \geq 0 : |X| \leq M \text{ a.s.}\})$$

All r.v.'s are defined on the same probability space (Ω, \mathcal{A}, P) .

Particular cases of interest

$p=1, E|X_n - X| \rightarrow 0$ (convergence in mean - σύγκλιση κατά μέσο)

$p=2, E(X_n - X)^2 \rightarrow 0$ (σύγκλιση κατά μέσο τετράγωνο - convergence in mean square)

$p=\infty, \|X_n - X\|_\infty \rightarrow 0$ (συγκλίνει σχεδόν ομοιόμορφα ή ομοιόμορφα με πιθανότητα 1 - converges almost uniformly)

$p=0, E|X_n - X|^0 = P(X_n \neq X) \rightarrow 0$ (converges in L^0)
 "
 $d_0(X_n, X)$

iii) $X_n \xrightarrow{P} X \iff \forall \epsilon > 0, P(|X_n - X| > \epsilon) \xrightarrow{n \rightarrow \infty} 0$ (δύο είναι κλάσες ισοδύναμων πιθανομετρικά (βαζοντας στα ω) από την ϵ -ζώνη)
 or equiv. $\forall \epsilon > 0, P(|X_n - X| \leq \epsilon) \xrightarrow{n \rightarrow \infty} 1$

(Since these properties are limit properties, we have:)

$$P(|X_n - X| > \varepsilon) = P(|X_n - X| \geq \varepsilon) \rightarrow 0$$

We say that (X_n) converges in probability (κατά πιθανότητα) to X or stochastic (στοχαστικά)

$$X_n \xrightarrow{c} X \Leftrightarrow \forall \varepsilon > 0 \quad \sum_{n \geq 1} P(|X_n - X| > \varepsilon) < +\infty$$

c : completely

$$\Leftrightarrow \lim_{n \rightarrow \infty} \sum_{k \geq n} P(|X_k - X| > \varepsilon) = 0$$

$$\left(\sum_{a_n \geq 0} a_n \in \mathbb{R} \Leftrightarrow \lim_{k \geq n} \sum_{k \geq n} a_k = 0, \text{ όριο ουσίας μηδενικό} \right)$$

Then we say that (X_n) converges completely to X . (it is used as a tool to deduce a.s. convergence)

Proposition (characterisation of $\xrightarrow{\text{a.s.}}$)

$$X_n \xrightarrow{\text{a.s.}} X \stackrel{(1)}{\Leftrightarrow} \forall \varepsilon > 0, P(\limsup_n |X_n - X| > \varepsilon) = 0$$

$$\stackrel{(2)}{\Leftrightarrow} \forall \varepsilon > 0, P(\underbrace{\sup_{k \geq n} |X_k - X|}_{Y_n} > \varepsilon) \rightarrow 0$$

$$\stackrel{(3)}{\Leftrightarrow} \sup_{k \geq n} |X_k - X| \xrightarrow{P} 0$$

$$(|X_n - X| \leq \sup_{k \geq n} |X_k - X|)$$

$$\stackrel{(4)}{\Leftrightarrow} \forall \varepsilon > 0 \quad P\left(\bigcup_{k \geq n} \{|X_k - X| > \varepsilon\}\right) \rightarrow 0$$

Proof

$$(1) X_n \xrightarrow{\text{a.s.}} X \iff |X_n - X| \xrightarrow{\text{a.s.}} 0 \iff \limsup_n |X_n - X| \stackrel{(*)}{=} 0$$

$$(*) \quad 0 \leq \liminf |X_n - X| \leq \limsup |X_n - X|. \quad \text{As } \limsup \stackrel{\text{a.s.}}{=} 0 \Rightarrow \liminf \stackrel{\text{a.s.}}{=} 0 \Rightarrow |X_n - X| \rightarrow 0.$$

$$\iff \forall \epsilon > 0, \limsup |X_n - X| \stackrel{\text{a.s.}}{\leq} \epsilon \quad (\text{also for } \epsilon = 1/n \quad \mathbb{1} \uparrow \forall \epsilon > 0 \text{ and } X \leq 0)$$

$$X \stackrel{\text{a.s.}}{\leq} 0 \Rightarrow X \stackrel{\text{a.s.}}{=} 0. \quad \text{Then for } \epsilon = 1/n, X \leq 1/n,$$

$$P(X > 0) = \lim_n \underbrace{P(X > \frac{1}{n})}_0 = 0$$

$$\iff \forall \epsilon > 0, P(\limsup |X_n - X| > \epsilon) = 0$$

(2) Let $\epsilon > 0$.

$$\{\limsup_n |X_n - X| > \epsilon\} = \{\omega \in \Omega : \limsup_n |X_n^{(\omega)} - X^{(\omega)}| > \epsilon\} \begin{matrix} \text{we work} \\ \iff \\ \text{with } \gamma \epsilon \end{matrix}$$

$$\limsup_n |X_n(\omega) - X(\omega)| > \epsilon \iff \liminf_n \sup_{k \geq n} |X_k(\omega) - X(\omega)| > \epsilon$$

(lim inf φ divouosa fe φω inf arnis)

$$\iff \forall n \geq 1 \quad \sup_{k \geq n} |X_k(\omega) - X(\omega)| > \epsilon$$

$$\iff \omega \in \bigcap_{n \geq 1} \left\{ \sup_{k \geq n} |X_k - X| > \epsilon \right\}.$$

So,

$$\{\limsup |X_n - X| \geq \varepsilon\} = \bigcap_{n \geq 1} \underbrace{\left\{ \sup_{k \geq n} |X_k - X| \geq \varepsilon \right\}}_{\text{decreasing seq. of events}}$$

We conclude that

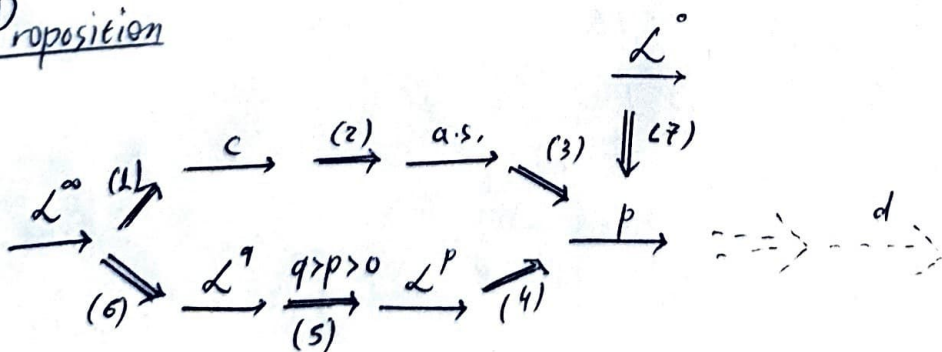
$$\begin{aligned} P(\limsup |X_n - X| \geq \varepsilon) &= P\left(\bigcap_{n \geq 1} \left\{ \sup_{k \geq n} |X_k - X| \geq \varepsilon \right\}\right) \\ &= P\left(\lim_n \left\{ \sup_{k \geq n} |X_k - X| \geq \varepsilon \right\}\right) \\ &= \lim_n P(\sup_{k \geq n} |X_k - X| \geq \varepsilon) \end{aligned}$$

(3) Let $Y_n = \sup_{k \geq n} |X_k - X|$, Then equivalently $Y_n \xrightarrow{p} 0$

(4) $\sup_{k \geq n} |X_k(\omega) - X(\omega)| \geq \varepsilon \iff \omega \in \bigcup_{k \geq n} \{|X_k - X| \geq \varepsilon\}$. So

$$\left\{ \sup_{k \geq n} |X_k - X| \geq \varepsilon \right\} = \bigcup_{k \geq n} \{|X_k - X| \geq \varepsilon\}, \text{ so } (2) \iff (4)$$

Proposition



Additionally, no implication holds as an equivalence. Without supplementary conditions.

Proof

$$(1) \xrightarrow{L^\infty} \rightarrow \xrightarrow{c}$$

Let $\varepsilon > 0$. Then

$$\exists n_0(\varepsilon): \forall n > n_0, \|X_n - X\|_\infty \leq \varepsilon.$$

Also, $|X_n - X| \stackrel{a.s.}{\leq} \|X_n - X\|_\infty$, so

$$\forall n > n_0 \quad |X_n - X| \stackrel{a.s.}{\leq} \varepsilon \rightarrow P(|X_n - X| > \varepsilon) = 0 \quad \forall n > n_0.$$

We deduce that:

$$\sum_{k > n_0} P(|X_k - X| > \varepsilon) = 0, \quad \forall n > n_0.$$

$$\Rightarrow \lim_{n \rightarrow \infty} \sum_{k > n} P(|X_k - X| > \varepsilon) = 0 \Rightarrow \sum_{k > n} P(|X_k - X| > \varepsilon) \stackrel{L^1}{\rightarrow} 0 \Rightarrow X_n \xrightarrow{c} X.$$

$$(2) \left(\xrightarrow{c} \rightarrow \xrightarrow{a.s.} \right)$$

Reminder: $X_n \xrightarrow{a.s.} X \Leftrightarrow \forall \varepsilon > 0 \quad P\left(\bigcup_{k > n} |X_k - X| > \varepsilon\right) \rightarrow 0.$

By assumption,

$$X_n \xrightarrow{c} X \rightarrow \lim_n \sum_{k > n} P(|X_k - X| > \varepsilon) = 0.$$

But

$$P\left(\bigcup_{k \geq n} |X_k - X| > \varepsilon\right) \stackrel{\sigma\text{-sub}}{\leq} \sum_{k \geq n} P(|X_k - X| > \varepsilon) \xrightarrow[n \rightarrow \infty]{\text{additivity}} 0$$

$$\Rightarrow X_n \xrightarrow{\text{a.s.}} X.$$

(3) Let $\varepsilon > 0$.

$$P(|X_n - X| > \varepsilon) \leq P\left(\sup_{k \geq n} |X_k - X| > \varepsilon\right) \xrightarrow[n \rightarrow \infty]{(*)} 0$$

(*) by assumption and the equiv. characterisation $X_n \xrightarrow{\text{a.s.}} X \Leftrightarrow \sup_{k \geq n} |X_k - X| \xrightarrow{P} 0$

($0 \leq R_n \leq S_n$ and $S_n \xrightarrow{P} 0$, then: $R_n \xrightarrow{P} 0$)

$$0 \leq |X_n - X| \leq \sup_{k \geq n} |X_k - X| \xrightarrow{P} 0$$

$$(4) \left(\xrightarrow{\mathcal{L}^p} \Rightarrow \xrightarrow{P} \right)$$

let $\varepsilon > 0$, $0 < p < \infty$

$$P(|X_n - X| > \varepsilon) = P(|X_n - X|^p > \varepsilon^p) \stackrel{\text{Markov}}{\leq} \frac{E|X_n - X|^p}{\varepsilon^p} \xrightarrow[n \rightarrow \infty]{\text{by assumption}} 0$$

$$\Rightarrow X_n \xrightarrow{P} X.$$

$$(5) \left(\xrightarrow{\mathcal{L}^q} \xrightarrow{q > p > 0} \xrightarrow{\mathcal{L}^p} \right)$$

$$\text{If } X_n \xrightarrow{\mathcal{L}^q} X \Rightarrow E|X_n - X|^q \rightarrow 0 \Rightarrow \|X_n - X\|_q \rightarrow 0 \xrightarrow{\text{Lyapunov}} \begin{matrix} \|Y\|_p \leq \|Y\|_q \\ (0 < p < q < \infty) \end{matrix}$$

$$\|X_n - X\|_p \rightarrow 0 \Rightarrow E|X_n - X|^p \rightarrow 0.$$

So $X_n \xrightarrow{L^p} X$ (επισημασμένο σε αυτήν τον χώρο από την ορισμένη, έχουν διάφορες πεπερασμένες φορές κλπ. Ούτως ούτως, ζήτησε η ακολουθία να προσεγγιστεί μέσα).

$$(6) \left(\xrightarrow{L^\infty} \Rightarrow \xrightarrow{L^q} \right)$$

$$0 \leq \|X_n - X\|_q \leq \|X_n - X\|_\infty.$$

$$\text{If } X_n \xrightarrow{L^\infty} X \Rightarrow \|X_n - X\|_\infty \rightarrow 0 \xrightarrow{\text{ineq}} \|X_n - X\|_q \rightarrow 0 \Rightarrow X_n \xrightarrow{L^q} X.$$

$$(7) \left(\xrightarrow{L^0} \Rightarrow \xrightarrow{P} \right).$$

Let $\varepsilon > 0$.

$$0 \leq P(X_n \neq X) \leq P(|X_n - X| > \varepsilon) \xrightarrow{\text{by assume}} 0 \Rightarrow X_n \xrightarrow{P} X$$

Counterexamples

$$(1) \xrightarrow{L^p} \not\Rightarrow \xrightarrow{\text{a.s.}}, p \in [0, +\infty), \text{ so}$$

$$\xrightarrow{P} \not\Rightarrow \xrightarrow{\text{a.s.}} \left(\begin{array}{l} \xrightarrow{L^p} \Rightarrow \xrightarrow{P} \\ \text{if } \xrightarrow{P} \Rightarrow \xrightarrow{\text{a.s.}} \end{array} \right) \xrightarrow{\text{we showed this}} \xrightarrow{L^p} \Rightarrow \xrightarrow{\text{a.s.}}$$

Let $(\Omega, \mathcal{A}, P) = ((0, 1), \mathcal{B}(0, 1), \mathcal{L})$ and set

$$X_n = \mathbf{1}_{A_n}$$

where (A_n) is the "triangular" sequence of partitions of $(0, 1)$ with successively decreasing length $\frac{1}{2}$.

We will show that

$$X_n \xrightarrow{L^p} X \Leftrightarrow E|X_n - X|^p \rightarrow 0, \text{ where } X = 0.$$

We have:

$$X_n \xrightarrow{L^p} 0 \Leftrightarrow E|X_n|^p \rightarrow 0$$

$$E|X_n|^p = E(\mathbb{1}_{A_n}^p) = E(\mathbb{1}_{A_n}) = P(A_n) \xrightarrow[n \rightarrow \infty]{\text{shown.}} 0.$$

> and since $\omega \in A$ just before n

But, $X_n \not\xrightarrow{\text{a.s.}} 0$, because $\limsup X_n = \limsup \mathbb{1}_{A_n} = 1 \quad \forall \omega \in (0, 1)$
 ($\mathbb{1}_{A_n}(\omega) = 1$ for infinite n).

Indeed $X_n \not\xrightarrow{\text{a.s.}} 0$.

$$(2) \xrightarrow{\text{a.s.}} \not\xrightarrow{L^p}, \quad 0 < p < +\infty \quad \text{and} \quad \xrightarrow{L^0} \not\xrightarrow{L^p}.$$

In the space of (L) ,

$$X_n = n \cdot \mathbb{1}_{(0, \frac{1}{n})}.$$

Obviously, $X_n(\omega) \xrightarrow[n \rightarrow \infty]{} 0, \quad \forall \omega \in (0, 1) \rightarrow X_n \xrightarrow{\text{a.s.}} 0.$

But

$$E(|X_n - 0|^p) = E(X_n^p) = E(n \cdot \mathbb{1}_{(0, \frac{1}{n})}) = n \cdot P((0, \frac{1}{n})) = n \cdot \frac{1}{n} = 1 \not\xrightarrow[n \rightarrow \infty]{} 0$$

$$\rightarrow X_n \not\xrightarrow{L^p} 0$$

But also: $P(X_n \neq 0) = \lambda(0, \frac{1}{n}) = \frac{1}{n} \rightarrow 0$.

So, $X_n \xrightarrow{L^0} 0$.

(3) $L^\infty \not\rightarrow L^0$, so $\xrightarrow{(\neq L^0)}$ $\not\rightarrow$ $\xrightarrow{L^0}$. *whatever convergence*

Take $X_n(\omega) = \frac{1}{n}, \forall \omega \in \Omega$.

$\|X_n - X\|_\infty \leq \sup_{\omega \in \Omega} |X_n(\omega) - 0| = \frac{1}{n} \rightarrow 0 \Rightarrow X_n \xrightarrow{L^\infty} X$
" *ess sup*

$P(X_n \neq 0) = P(\frac{1}{n} \neq 0) = 1 \not\rightarrow 0$.

So $X_n \not\xrightarrow{L^0} X$

Assume $\xrightarrow{(\neq L^0)} \Rightarrow \xrightarrow{L^0}$. Then $\xrightarrow{L^\infty} \Rightarrow \xrightarrow{(\neq L^0)}$. So

$\xrightarrow{L^\infty} \Rightarrow \xrightarrow{L^0}$ impossible contradiction.