

## SEs and 95% CI for linear combination of the estimates in STATA

In Stata we can estimate the mean, the SE and the 95% CI of a linear combination of the parameters by using the command `lincom` after the model fit.

For example, to estimate mean plasma of retinol in men with infrequent vitamin use we need to estimate the combination:  $Comb1 = b_{men} + b_{infreq} + b_{men\_infreq} + \_cons$ .

If the model has been defined as: `reg retplasm men freq infeq menfreq meninfr`

We can get the estimate, the SE and the 95% CI of the parameter `Comb1` in Stata by the command:

```
Lincom men+infreq+men_infreq+_cons
```

after fitting the model.

If the model has been defined using the `xi` command, we must trace and use the appropriate dummy variables (`I*`) in Stata's "variables" list.

For example:

```
. char sex[omit] 2
. char vituse[omit] 3
. xi:reg  retplasm i.sex*i.vituse
i.sex      _Isex_1-2      (naturally coded; _Isex_2 omitted)
i.vituse   _Ivituse_1-3  (naturally coded; _Ivituse_3 omitted)
i.sex*i.vituse  _IsexXvit_#_#  (coded as above)
```

Source	SS	df	MS	Number of obs	=	314
-----+-----						
				F(5, 308)	=	4.03
Model	836388.279	5	167277.656	Prob > F	=	0.0015
Residual	12783793.6	308	41505.8233	R-squared	=	0.0614
-----+-----						
				Adj R-squared	=	0.0462
Total	13620181.9	313	43514.958	Root MSE	=	203.73

  

retplasm	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
_Isex_1	166.3468	47.76693	3.48	0.001	72.35603	260.3376
_Ivituse_1	33.46968	29.28935	1.14	0.254	-24.16285	91.10221
_Ivituse_2	39.49589	31.87656	1.24	0.216	-23.22749	102.2193
_IsexXvit_1_1	-11.72721	76.51943	-0.15	0.878	-162.2942	138.8398
_IsexXvit_1_2	-255.6611	105.4603	-2.42	0.016	-463.175	-48.14725
_cons	563.2184	21.84213	25.79	0.000	520.2397	606.1971
-----+-----						

Stata has now created a dummy variable for sex (`Isex_1`), two dummy variables for vitamine use (`Ivituse_1` `Ivituse_2`) and two dummy variables for their respective interactions (`IsexXvit_1_1` `IsexXvit_1_2`). Notice the reference levels for these two categorical variables as defined by the “`char varname[omit] #`” commands. Thus if we want to estimate mean retinol plasma levels for men who use vitamins infrequently (`sex=2 & vituse=2`) we can use the following command:

```
. lincom Issex_1+ Ivituse_2+ IssexXvit_1_2+_cons

( 1)  Issex_1 + Ivituse_2 + IssexXvit_1_2 + _cons = 0.0
```

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
(1)	513.4	91.11073	5.635	0.000	334.1218	692.6782

These results are based on standard probability theory [ $\text{Var}(A+B)=\text{Var}(A)+\text{Var}(B)+2*\text{Cov}(A, B)$ ], and can be obtained “by hand” using some simple matrix algebra and taking advantage of the fact that Stata saves the vector of coefficients as `e(b)` and their respective variance-covariance matrix as `e(V)` after each model estimation command:

We can print these two matrices using the “`matrix list`” command.

```
. mat li e(b)
e(b) [1,6]
      _Isex_1  _Ivituse_1  _Ivituse_2  _IsexXvit_~1  _IsexXvit_~2  _cons
y1    166.34683  33.469683  39.495895  -11.727208  -255.66111  563.21839

. mat li e(V)
symmetric e(V) [6,6]
      _Isex_1  _Ivituse_1  _Ivituse_2  _IsexXvit_~1  _IsexXvit_~2  _cons
_Isex_1    2281.6794
_Ivituse_1  477.07843  857.8658
_Ivituse_2  477.07843  477.07843  1016.1151
_IsexXvit_~1 -2281.6794  -857.8658  -477.07843  5855.2225
_IsexXvit_~2 -2281.6794  -477.07843  -1016.1151  2281.6794  11121.881
_cons      -477.07843  -477.07843  -477.07843  477.07843  477.07843  477.07843
```

Check the relation between the output of the `regress` command and the elements of those two matrices.

We can now estimate any linear combination of the coefficients of the model and its standard error (or variance). The calculations can be simplified by defining a vector `C` of length 6 ( our model has 6 coefficients) with all its elements equal to 0 except those who correspond to the coefficients that we

want to include in the linear combinations (constraint). These elements must be set equal to 1. For example if we want to replicate the `lincom` command as presented above our `C` vector must be `C=(1, 0, 1, 0, 1, 1)`.

```
. matrix C=(1,0,1,0,1,1)
```

By multiplying the `1x6 e(b)` matrix with the transpose of the `1x6 C` vector we have a scalar (`1x1`) which is equal with the sum of the appropriate model coefficients

```
(Isex_1 + Ivituse_1 + Ivituse_2 + IsexXvit_1_1 + IsexXvit_1_2 + _cons)
```

```
. matrix comb1=e(b)*C'  
. mat li comb1  
symmetric comb1[1,1]  
      r1  
y1  513.4
```

The variance (and the standard error of this estimate) can be easily obtained by the following matrix product  $\text{Var}(\text{comb1})=C*e(V)*C^T$

```
. matrix Vcomb1=C*e(V)*C'  
. matrix li Vcomb1  
  
symmetric Vcomb1[1,1]  
      r1  
r1  8301.1647
```

The standard error of `comb1` is just the square root of its variance and can be obtained by the following two ways with the second being more general:

```
. local vcomb1=Vcomb1[1,1]  
. di sqrt(`vcomb1')  
91.110727
```

OR

```
. mat SEcomb1=cholesky(Vcomb1)  
. mat li SEcomb1  
symmetric SEcomb1[1,1]  
      r1  
r1  91.110727
```

Notice that in the first approach we used Stata's local macros to extract the element of `Vcomb1` as a scalar in order to calculate its square root, while in the second we used the Cholesky matrix decomposition which in the 1-dimensional case is equivalent to the square root. The second approach is more general while it allows for multiple linear combinations' estimation along with their standard errors by just replacing the `1xp C` vector (`p= # of coefficients in the model`) with the appropriate `m xp`

matrix where  $m$  is the number of linear combinations to be estimated. The  $C$  matrix must be formed in the same manner as the  $C$  vector in our example\*.

The 95% confidence interval can be obtained by adding or subtracting the product of the standard error with the  $t_{0.975,df}$  value (remember that we have an OLS estimation case):

```
. *Lower 95% Limit
. di 513.4 -invttail(308,0.025)*(91.11073)
334.12179
. *Upper 95% Limit
. di 513.4 +invttail(308,0.025)*(91.11073)
692.67821
```

Notice the number of degrees of freedom ( $df=308$ ) and the  $p$  argument (0.025) in the `invttail(df,p)` Stata's function. In case of ML estimation we must use the  $z_{0.975}$  value instead of the  $t_{0.975,df}$  one.

The procedure described above is a general method for estimating linear combinations of a model's coefficients and can be very useful especially in statistical packages lacking the "lincom" Stata's utility.

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\*We can estimate two linear combinations (men infrequent and men frequent users) simultaneously as follows:

```
. matrix C=(1,0,1,0,1,1\1,1,0,1,0,1)
. matrix comb1=e(b)*C'
. mat li comb1
comb1[1,2]
      r1      r2
y1      513.4 751.30769
. matrix Vcomb1=C*e(V)*C'
. di sqrt(Vcomb1[1,1])
91.110727
. di sqrt(Vcomb1[2,2])
56.504475
```

Compare with the following `lincom` commands:

```
. lincom _Isex_1+ _Ivituse_2+ _IsexXvit_1_2+ _cons
( 1)  _Isex_1 + _Ivituse_2 + _IsexXvit_1_2 + _cons = 0
```

retplasm	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
(1)	513.4	91.11073	5.63	0.000	334.1218 692.6782

```
. lincom _Isex_1+ _Ivituse_1+ _IsexXvit_1_1+ _cons
( 1)  _Isex_1 + _Ivituse_1 + _IsexXvit_1_1 + _cons = 0
```

retplasm	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
(1)	751.3077	56.50447	13.30	0.000	640.1241 862.4913