

## Notes for laboratory session 9

Produce two 2x2 tables corresponding to the association between "Mammograph experience" and "Family history". In the first table use only the "Within one year" and "Never" categories while in the second use the "Over a year ago" and "Never" categories of the "Mammograph experience" variable.

```
. tab ME hist if ME==0 | ME==1

      Mammograph |      Fam. history
      experience |          No       Yes |      Total
-----+-----+-----+
      Never |      220        14 |      234
  Within one year |      85        19 |      104
-----+-----+-----+
      Total |      305        33 |      338

. tab ME hist if ME==0 | ME==2

      Mammograph |      Fam. history
      experience |          No       Yes |      Total
-----+-----+-----+
      Never |      220        14 |      234
  Over a year ago |      63        11 |      74
-----+-----+-----+
      Total |      283        25 |      308
```

a) Calculate the Odds Ratios associated with the tables above.

Now fit the logistic regression models corresponding to the previous tables .

```
. xi: logit ME i.hist if ME==1 | ME==0, nolog
i.hist           Ihist_0-1      (naturally coded; Ihist_0 omitted)

Logit estimates                                         Number of obs =      338
                                                       LR chi2(1) =     11.32
                                                       Prob > chi2 =  0.0008
Log likelihood = -202.96528                           Pseudo R2 =    0.0271

-----+
      ME |      Coef.      Std. Err.      z      P>|z|      [95% Conf. Interval]
-----+
      Ihist_1 |   1.256358   .3746603   3.353   0.001      .5220373   1.990679
      _cons |  -.9509763   .1277112  -7.446   0.000     -1.201286  -.7006669
-----+
. xi: logit ME i.hist if ME==2 | ME==0, nolog
i.hist           Ihist_0-1      (naturally coded; Ihist_0 omitted)

Logit estimates                                         Number of obs =      308
                                                       LR chi2(1) =     5.26
                                                       Prob > chi2 =  0.0218
Log likelihood = -167.19417                           Pseudo R2 =    0.0155

-----+
      ME |      Coef.      Std. Err.      z      P>|z|      [95% Conf. Interval]
-----+
      Ihist_1 |   1.009331   .4274999   2.361   0.018      .1714464   1.847215
      _cons |  -1.250493   .1428932  -8.751   0.000     -1.530558  -.9704273
-----+
```

Now perform the above two logistic regressions in one step using the multinomial logit (mlogit) command of STATA :

```
. xi: mlogit ME i.hist, nolog
i.hist          Ihist_0-1      (naturally coded; Ihist_0 omitted)

Multinomial regression
Number of obs     =        412
LR chi2(2)       =      12.86
Prob > chi2      =    0.0016
Pseudo R2        =    0.0160

Log likelihood = -396.16997
-----+
ME |      Coef.    Std. Err.      z     P>|z|      [95% Conf. Interval]
-----+
Within o |
Ihist_1 |   1.256358   .3746603     3.353   0.001     .5220372   1.990679
_cons |  -.9509763   .1277112    -7.446   0.000    -1.201286  -.7006669
-----+
Over a y |
Ihist_1 |   1.009331   .4274998     2.361   0.018     .1714466   1.847215
_cons |  -1.250493   .1428932    -8.751   0.000    -1.530558  -.9704273
-----+
(Outcome ME==Never is the comparison group)

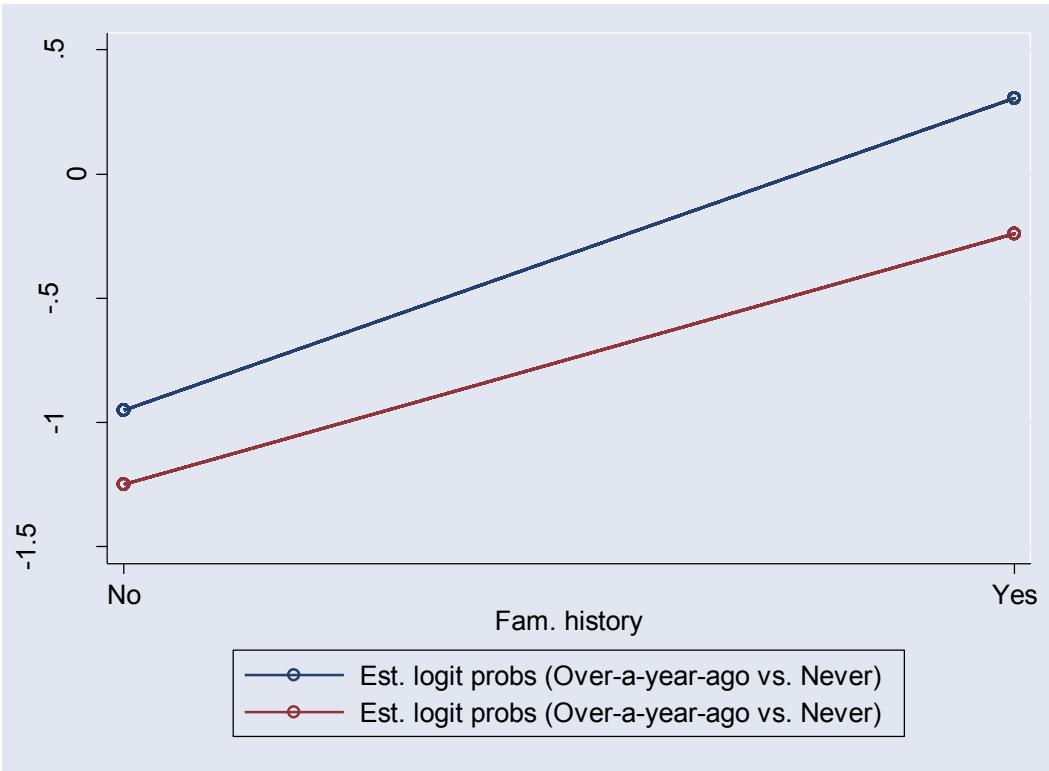
. est store A1
```

- b) Compare the output of the “mlogit” command to the output of the two “logit” commands.
- c) What is the interpretation of the coefficients in the “mlogit” command output. Compare with the results of question a).

## Graphical inspection of the results

We can inspect graphically the results as follows:

```
. quietly xi: mlogit ME i.hist
. predict p1 p2 p3
(option p assumed; predicted probabilities)
. gen logit_1=log(p2/p1)
. label var logit_1 "Est. logit probs (Within-one-year vs. Never)"
. gen logit_2=log(p3/p1)
. label var logit_2 "Est. logit probs (Over-a-year-ago vs. Never)"
. sc logit_1 logit_2 hist, c(1 1) xlab(0 1, val) legend(row(2)) ylab()
yti("Estimated logits")
```



## Testing the equality of the two odds ratios

Test whether  $OR_1=OR_2$  by testing whether the odds ratio that corresponds to the following table is one:

```
. tab ME hist if ME==1 | ME==2,chi

      Mammograph |      Fam. history
      experience |          No       Yes |      Total
-----+-----+-----+
  Within one year |        85       19 |       104
  Over a year ago |        63       11 |       74
-----+-----+-----+
      Total |        148       30 |       178

Pearson chi2(1) =    0.3576   Pr = 0.550
```

- d) Calculate the OR and its SE in order to produce the relevant statistic and test whether  $OR=1$ .

We can use this alternative approach to test the same hypothesis ( $OR_1=OR_2$ ) :

```
. test [1]
(1) [Within o]Ihist_1 = 0.0
      chi2( 1) =    11.24
      Prob > chi2 =    0.0008

. test [2]
(1) [Over a y]Ihist_1 = 0.0
      chi2( 1) =     5.57
      Prob > chi2 =    0.0182

. test [1=2]
(1) [Within o]Ihist_1 - [Over a y]Ihist_1 = 0.0
      chi2( 1) =     0.36
      Prob > chi2 =    0.5505
```

- e) Notice the relation between the chi-square statistics in the first two "test" commands and the z statistics in the previous "mlogit" command.

Now fit the null model in order to check the significance of "family history" by using a likelihood-ratio test approach.

```
. mlogit ME, nolog
Multinomial regression
Number of obs      =        412
LR chi2(0)        =       0.00
Prob > chi2       =
Pseudo R2         =       0.0000

Log likelihood = -402.59901

-----+
          ME |      Coef.      Std. Err.          z      P>|z|
-----+
Within o |
      _cons |   -.8109302   .1178511      -6.881    0.000      -1.041914   -.5799462
-----+
Over a y |
      _cons |   -1.151256   .133368      -8.632    0.000      -1.412652   -.8898596
-----+
(Outcome ME==Never is the comparison group)

. est store A0
```

```
. lrtest A0 A1  
likelihood-ratio test  
(Assumption: A0 nested in A1) LR chi2(2) = 12.86  
Prob > chi2 = 0.0016
```

f) Calculate the likelihood-ratio test manually.

The significance of "family history" can also be checked by a Wald type test:

```
. quietly xi: mlogit ME i.hist, nolog  
. test Ihist_1  
( 1) [Within o]Ihist_1 = 0.0  
( 2) [Over a y]Ihist_1 = 0.0  
chi2( 2) = 12.01  
Prob > chi2 = 0.0025
```

g) What is the conclusion of the above tests?

## Incorporating a polytomous covariate

Tabulate the data according to the variables "mammograph experience" and "detc"

```
. tab ME detc
```

Mammograph experience	Likely find cancer			Total
	Not likel	Somewhat	Very like	
Never	13	77	144	234
Within one year	1	12	91	104
Over a year ago	4	16	54	74
Total	18	105	289	412

Now define the level 1 of the factor "detc" as reference category and then use the "mlogit" STATA command in order to investigate the association between these two variables.

```
. char detc[omit] 1
. xi: mlogit ME i.detc, nolog
i.detc           Idetc_1-3      (naturally coded; Idetc_1 omitted)

Multinomial regression                                         Number of obs = 412
                                                               LR chi2(4) = 26.80
                                                               Prob > chi2 = 0.0000
                                                               Pseudo R2 = 0.0333

Log likelihood = -389.20054
```

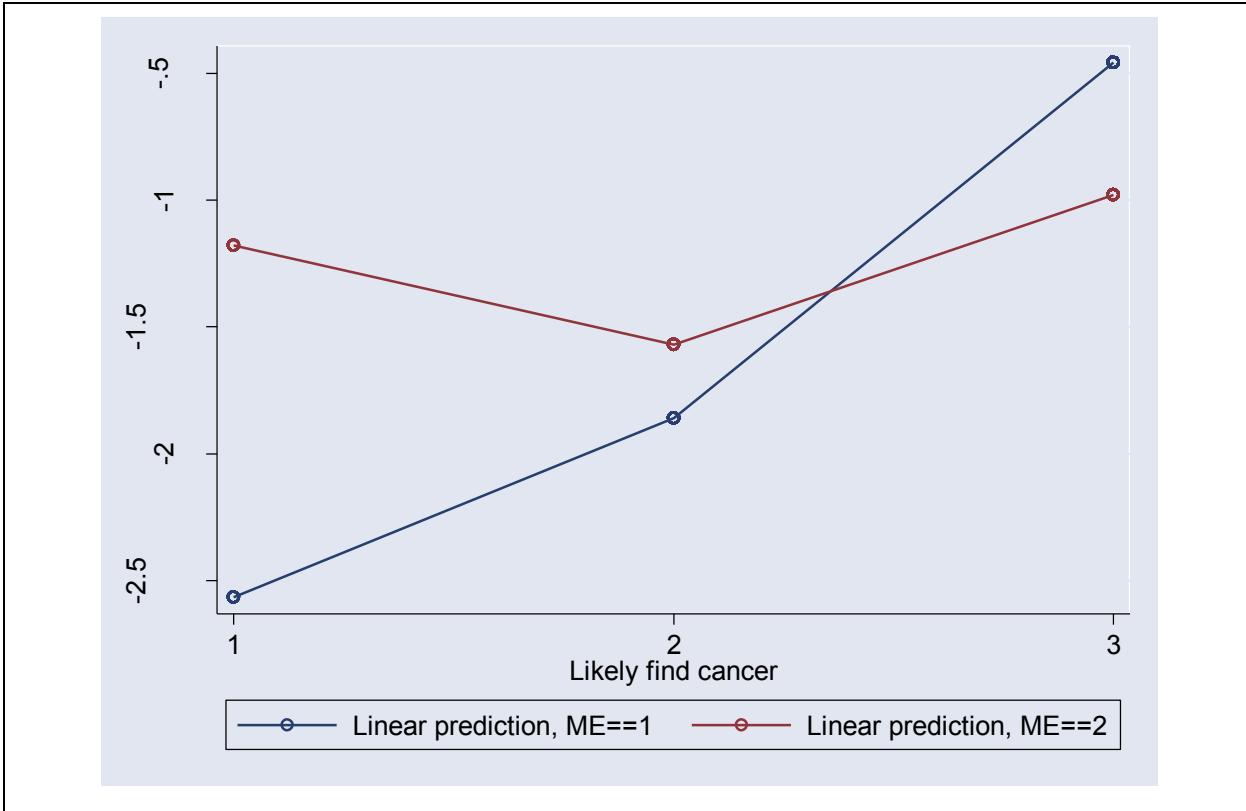
ME	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
Within o					
Idetc_2	.7060506	1.083136	0.652	0.514	-1.416856 2.828958
Idetc_3	2.105996	1.046325	2.013	0.044	.0552361 4.156755
_cons	-2.564949	1.03772	-2.472	0.013	-4.598843 -.5310556
Over a y					
Idetc_2	-.3925617	.6343589	-0.619	0.536	-1.635882 .850759
Idetc_3	.1978257	.5936221	0.333	0.739	-.9656522 1.361304
_cons	-1.178655	.5717729	-2.061	0.039	-2.299309 -.0580007

(Outcome ME==Never is the comparison group)

h) Try to interpret the coefficients of the Idetc\_2 and Idetc\_3 dummy variables in terms of ORs produced by appropriate sub-tables of the 3x3 table above.

### Graphical inspection of the model

```
sort detc
cap drop logit_1 logit_2
predict logit_1,xb outcome(1)
predict logit_2,xb outcome(2)
sc logit_1 logit_2 detc, c(l l) xlab(1 2 3) ylab() l1("log odds")
```



We can now use the "test" command for hypotheses testing

```
. test [1]
( 1) [Within o]Idetc_2 = 0.0
( 2) [Within o]Idetc_3 = 0.0

chi2( 2) =     20.41
Prob > chi2 =      0.0000

. test [2]
( 1) [Over a y]Idetc_2 = 0.0
( 2) [Over a y]Idetc_3 = 0.0

chi2( 2) =      3.46
Prob > chi2 =      0.1773

. test [1=2]
( 1) [Within o]Idetc_2 - [Over a y]Idetc_2 = 0.0
( 2) [Within o]Idetc_3 - [Over a y]Idetc_3 = 0.0

chi2( 2) =      6.20
Prob > chi2 =      0.0450
```

- i) State the null hypotheses for the three "test" commands.

### Assessment of the significance of a continuous factor

Fit the multinomial logistic regression model for the effect of perceived benefit (pb) of mammography (higher scores denote a smaller perceived benefit).

```
. mlogit ME pb,nolog

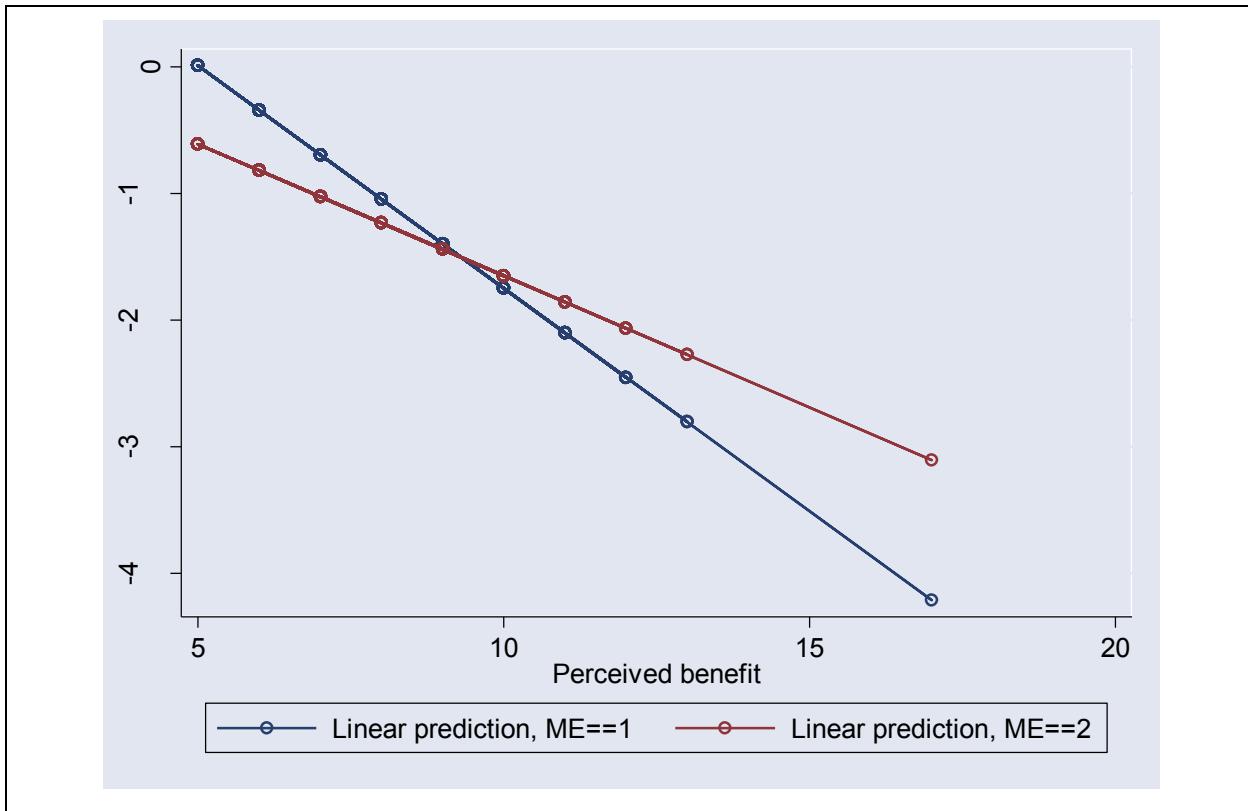
Multinomial regression                                         Number of obs      =      412
Log likelihood = -384.97236                                     LR chi2(2)        =     35.25
                                                               Prob > chi2       =  0.0000
                                                               Pseudo R2        =  0.0438

-----+
          ME |      Coef.    Std. Err.      z     P>|z|   [95% Conf. Interval]
-----+
Within o |
    pb |   -.351685   .0666009    -5.280    0.000    -.4822205   -.2211496
    _cons |   1.76865   .484798     3.648    0.000     .8184631   2.718836
-----+
Over a y |
    pb |   -.207975   .0684675    -3.038    0.002    -.3421688   -.0737812
    _cons |   .4314007   .5228924     0.825    0.409    -.5934496   1.456251
-----+
(Outcome ME==Never is the comparison group)
```

*k) What is the interpretation of the “pb” coefficients?*

We can see graphically the model as follows:

```
. predict lhat1, xb outcome(1)
. predict lhat2, xb outcome(2)
. sc lhat* pb, xlab() ylab() c(l l)
```



**Testing the effect of a continuous covariate**

The overall effect of the perceived benefit on the likelihood of a mammogram is tested as follows:

```
. test pb  
( 1) [Within o]pb = 0.0  
( 2) [Over a y]pb = 0.0  
  
chi2( 2) =    31.00  
Prob > chi2 =    0.0000
```

On the other hand, we can test whether the relationship between pb and ME is the same for both levels 1 and 2 compared to level 0 :

```
. test [1]pb=[2]pb  
( 1) [Within o]pb - [Over a y]pb = 0.0  
  
chi2( 1) =    3.02  
Prob > chi2 =    0.0821
```

- l) Interpret the result of the previous test command and compare it with the graphical representation of the model.*

### The method of Begg & Gray (Biometrika, 1984)

Now use the method of Begg & Gray by fitting two logistic regression models instead of one multinomial logistic regression model

. xi: logit ME i.sympt pb i.hist BSE i.detc if ME==1 ME==0, nolog					
				Pseudo R2	= 0.2324
ME	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
Isympt_2	.2333352	.9354078	0.249	0.803	-1.60003 2.066701
Isympt_3	2.003157	.7897107	2.537	0.011	.4553526 3.550961
Isympt_4	2.526276	.7863638	3.213	0.001	.9850313 4.067521
pb	-.2127205	.0767566	-2.771	0.006	-.3631606 -.0622804
Ihist_1	1.415273	.4687315	3.019	0.003	.4965762 2.33397
BSE	1.400521	.5382657	2.602	0.009	.34554 2.455503
Idetc_2	.20745	1.166733	0.178	0.859	-2.079304 2.494204
Idetc_3	1.050927	1.126617	0.933	0.351	-1.157202 3.259057
_cons	-3.376979	1.540057	-2.193	0.028	-6.395435 -.3585223

. xi: logit ME i.sympt pb i.hist BSE i.detc if ME==2 ME==0, nolog					
				Pseudo R2	= 0.1040
ME	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
Isympt_2	-.2744739	.6472962	-0.424	0.672	-1.543151 .9942034
Isympt_3	.8085675	.5421184	1.491	0.136	-.2539651 1.8711
Isympt_4	1.209314	.5479505	2.207	0.027	.135351 2.283278
pb	-.1426345	.075692	-1.884	0.060	-.2909881 .0057191
Ihist_1	1.156508	.4735987	2.442	0.015	.2282712 2.084744
BSE	1.018466	.5157325	1.975	0.048	.0076488 2.029283
Idetc_2	-.778484	.7179777	-1.084	0.278	-2.185694 .6287264
Idetc_3	-.566554	.686941	-0.825	0.410	-1.912934 .7798256
_cons	-1.156001	1.112774	-1.039	0.299	-3.336999 1.024996

The full multinomial logistic regression is as follows:

```
. xi: mlogit ME i.sympt pb i.hist BSE i.detc , nolog
i.sympt           Isympt_1-4      (naturally coded; Isympt_1 omitted)
i.hist            Ihist_0-1      (naturally coded; Ihist_0 omitted)
i.detc           Idetc_1-3      (naturally coded; Idetc_1 omitted)

Multinomial regression                                         Number of obs = 412
                                                               LR chi2(16) = 111.30
                                                               Prob > chi2 = 0.0000
Log likelihood = -346.95096                                     Pseudo R2 = 0.1382

-----+
          ME |      Coef.    Std. Err.      z     P>|z| [95% Conf. Interval]
-----+
Within o |
Isympt_2 |   .1100372   .9227608   0.119   0.905   -1.698541   1.918615
Isympt_3 |   1.924708   .7775975   2.475   0.013   .4006449   3.448771
Isympt_4 |   2.456993   .7753323   3.169   0.002   .9373693   3.976616
      pb |  -.2194368   .0755139  -2.906   0.004   -.3674414  -.0714323
     Ihist_1 |   1.366239   .4375196   3.123   0.002   .5087163   2.223762
      BSE |   1.291666   .529891    2.438   0.015   .2530992   2.330234
Idetc_2 |   .0170207   1.161896   0.015   0.988   -2.260254   2.294296
Idetc_3 |   .9041379   1.126822   0.802   0.422   -1.304393   3.112668
      _cons |  -2.99875   1.53922  -1.948   0.051   -6.015566   .0180663
-----+
Over a y |
Isympt_2 |  -.2900833   .6440636  -0.450   0.652   -1.552425   .9722582
Isympt_3 |   .8173136   .5397922   1.514   0.130   -.2406596   1.875287
Isympt_4 |   1.132239   .5476704   2.067   0.039   .0588252   2.205654
      pb |  -.1482068   .0763686  -1.941   0.052   -.2978866   .0014729
     Ihist_1 |   1.065436   .459396    2.319   0.020   .1650366   1.965836
      BSE |   1.052144   .5149894   2.043   0.041   .0427838   2.061505
Idetc_2 |  -.9243928   .7137382  -1.295   0.195   -2.323294   .4745083
Idetc_3 |  -.6905329   .6871078  -1.005   0.315   -2.037239   .6561736
      _cons |  -.9860915   1.111832  -0.887   0.375   -3.165242   1.193059
-----+
(Outcome ME==Never is the comparison group)
```

*m) Compare the results of the above two approaches.*